

**MSO201a: Probability and Statistics**  
**2019-20-II Semester**  
**Assignment No. 6**  
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1. Let  $X_1, X_2, \dots$  be a sequence of r.v.s, such that  $X_n, n = 1, 2, \dots$ , has the d.f.:  $F_n(x) = 0$ , if  $x < -n$ ,  $= \frac{x+n}{2n}$ , if  $-n \leq x < n$ , and  $= 1$ , if  $x \geq n$ . Does  $F_n(\cdot)$  converge to a d.f., as  $n \rightarrow \infty$ ?
2. Let  $X_1, X_2, \dots$  be a sequence of i.i.d. r.v.s and let  $X_{1:n} = \min\{X_1, \dots, X_n\}$  and  $Y_n = nX_{1:n}, n = 1, 2, \dots$ . Find the limiting distributions of  $X_{1:n}$  and  $Y_n$  (as  $n \rightarrow \infty$ ) when (a)  $X_1 \sim U(0, \theta), \theta > 0$ ; (b)  $X_1 \sim \text{Exp}(\theta), \theta > 0$ .
3. Let  $X_1, X_2, \dots$  be a sequence of independent r.v.s with  $P(X_n = x) = \frac{1}{2}$ , if  $x = -n^{\frac{1}{4}}, n^{\frac{1}{4}}$ , and  $= 0$ , otherwise. Show that  $\bar{X}_n \xrightarrow{P} 0$ , as  $n \rightarrow \infty$ .
4. (a) If  $X_n \xrightarrow{P} a$  and  $X_n \xrightarrow{P} b$ , then show that  $a = b$ .  
 (b) Let  $a$  and  $r > 0$  be real numbers. If  $E(|X_n - a|^r) \rightarrow 0$ , as  $n \rightarrow \infty$ , then show that  $X_n \xrightarrow{P} a$ .
5. (a) For  $r > 0$  and  $t > 0$ , show that  $E\left(\frac{|X|^r}{1+|X|^r}\right) - \frac{t^r}{1+t^r} \leq P(|X| \geq t) \leq \frac{1+t^r}{t^r} E\left(\frac{|X|^r}{1+|X|^r}\right)$ .  
 (b) Show that  $X_n \xrightarrow{P} 0 \Leftrightarrow E\left(\frac{|X_n|^r}{1+|X_n|^r}\right) \rightarrow 0$ , for some  $r > 0$ .
6. (a) If  $\{X_n\}_{n \geq 1}$  are identically distributed and  $a_n \rightarrow 0$ , then show that  $a_n X_n \xrightarrow{P} 0$ .  
 (b) If  $Y_n \leq X_n \leq Z_n, n = 1, 2, \dots, Y_n \xrightarrow{P} a$  and  $Z_n \xrightarrow{P} a$ , then show that  $X_n \xrightarrow{P} a$ .  
 (c) If  $X_n \xrightarrow{P} c$  and  $a_n \rightarrow a$ , then show that, as  $n \rightarrow \infty, X_n + a_n \xrightarrow{P} c + a$  and  $a_n X_n \xrightarrow{P} ac$ .  
 (d) Let  $X_n = \min(|Y_n|, a), n = 1, 2, \dots$ , where  $a$  is a positive constant. Show that  $X_n \xrightarrow{P} 0 \Leftrightarrow Y_n \xrightarrow{P} 0$ .
7. Let  $X_1, X_2, \dots$  be a sequence of i.i.d. r.v.s with mean  $\mu$  and finite variance. Show that:  
 (a)  $\frac{2}{n(n+1)} \sum_{i=1}^n iX_i \xrightarrow{P} \mu$ ;  
 (b)  $\frac{6}{n(n+1)(2n+1)} \sum_{i=1}^n i^2 X_i \xrightarrow{P} \mu$ .
8. Let  $X_n, n = 1, 2, \dots$ , have a negative binomial distribution with parameters  $n$  and  $p_n = 1 - \frac{\theta}{n}$ , i.e.,  $X_n$  has the p.m.f.  $P(X_n = x) = \binom{n+x-1}{x} p_n^n (1-p_n)^x, x = 0, 1, 2, \dots; n = 1, 2, \dots$ . Show that  $X_n \xrightarrow{d} X \sim \text{Poisson}(\theta)$ .

9. (a) Let  $X_n \sim \text{Gamma}(\frac{1}{n}, n)$ ,  $n = 1, 2, \dots$ . Show that  $X_n \xrightarrow{P} 1$ .  
 (b) Let  $X_n \sim N(\frac{1}{n}, 1 - \frac{1}{n})$ ,  $n = 1, 2, \dots$ . Show that  $X_n \xrightarrow{d} Z \sim N(0, 1)$ .
10. (a) Let  $f(x) = \frac{1}{x^2}$ , if  $1 \leq x < \infty$ , and  $= 0$ , elsewhere, be the p.d.f. of a r.v.  $X$ . Consider the random sample of size 72 from the distribution having p.d.f.  $f(\cdot)$ . Compute, approximately, the possibility that more than 50 of the items of the random sample are less than 3.  
 (b) Let  $X_1, X_2, \dots$  be a random sample from Poisson(3) distribution and let  $Y = \sum_{i=1}^{100} X_i$ . Find, approximately,  $P(100 \leq Y \leq 200)$ .  
 (c) Let  $X \sim \text{Bin}(25, 0.6)$ . Find, approximately,  $P(10 \leq X \leq 16)$ . What is the exact value of this probability?
11. (a) Show that  $\lim_{n \rightarrow \infty} e^{-n} \sum_{k=0}^n \frac{n^k}{k!} = \frac{1}{2}$ .  
 (b) Show that  $\lim_{n \rightarrow \infty} 2^{-n} \sum_{k=0}^{r_n} \binom{n}{k} = \frac{1}{2}$ , where  $r_n$  is the largest integer  $\leq \frac{n}{2}$ .
12. (a) If  $T_n = \max(|X_1|, \dots, |X_n|) \xrightarrow{P} 0$ , as  $n \rightarrow \infty$ , then show that  $\bar{X}_n \xrightarrow{P} 0$ . Is the conclusion true if only  $S_n = \max(X_1, \dots, X_n) \xrightarrow{P} 0$ .  
 (b) If  $\{X_n\}_{n \geq 1}$  are i.i.d.  $U(0, 1)$  r.v.s. and  $Z_n = (\prod_{i=1}^n X_i)^{\frac{1}{n}}$ ,  $n = 1, 2, \dots$ . Find a real  $\alpha$  such that  $Z_n \xrightarrow{P} \alpha$ .
13. Let  $\{E_n\}_{n \geq 1}$  be a sequence of i.i.d.  $\text{Exp}(1)$  r.v.s.  
 (a) Show that  $T_n \equiv \sum_{i=1}^n E_i \sim \text{Gamma}(n, 1)$ ,  $n = 1, 2, \dots$ .  
 (b) For any real number  $x$ , show that  $\lim_{n \rightarrow \infty} \int_0^{n+x\sqrt{n}} \frac{e^{-t} t^{n-1}}{\Gamma(n)} dt = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt$ .  
 (c) For large values of  $n$ , show that an approximation (called the Stirling approximation) to the gamma function is:  $\Gamma n \approx \sqrt{2\pi} e^{-n} n^{n-\frac{1}{2}}$ .
14. Let  $X_1, X_2, \dots$  be a sequence of i.i.d. r.v.s having the common Cauchy p.d.f.  $f(x) = \frac{1}{\pi} \cdot \frac{1}{1+x^2}$ ,  $-\infty < x < \infty$ .  
 (a) For any  $\alpha \in (0, 1)$ , show that  $Y = \alpha X_1 + (1 - \alpha)X_2$  again has a Cauchy p.d.f.  $f(\cdot)$ .  
 (b) Note that  $\bar{X}_{n+1} = \frac{n}{n+1} \bar{X}_n + \frac{1}{n+1} X_{n+1}$  and hence, using induction, conclude that  $\bar{X}_n$  has the same distribution as  $X_1$ .  
 (c) Show that  $\bar{X}_n$  does not converge in probability to any constant. (Note that  $E(X_1)$  does not exist and hence the WLLN is not guaranteed).
15. Let  $X_n \sim \text{Poisson}(4n)$ ,  $n = 1, 2, \dots$ , and let  $Y_n = \frac{X_n}{n}$ ,  $n = 1, 2, \dots$ .  
 (a) Show that  $Y_n \xrightarrow{P} 4$ ;  
 (b) Show that  $Y_n^2 + \sqrt{Y_n} \xrightarrow{P} 18$ ;  
 (c) Show that  $\frac{n^2 Y_n^2 + n Y_n}{n Y_n + n^2} \xrightarrow{P} 16$ .
16. Let  $\bar{X}_n$  be the sample mean computed from a random sample of size  $n$  from a distribution with mean  $\mu$  ( $-\infty < \mu < \infty$ ) and variance  $\sigma^2$  ( $0 < \sigma < \infty$ ). Let  $Z_n = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}$ .  
 (a) If  $Y_n \xrightarrow{P} 4$ , show that:  $\frac{4Z_n}{Y_n} \xrightarrow{d} Z \sim N(0, 1)$ ;  $\frac{16Z_n^2}{Y_n^2} \xrightarrow{d} U \sim \chi_1^2$ ; and  $\frac{(4n+Y_n)Z_n}{(nY_n+Y_n^2)} \xrightarrow{d}$

$Z \sim N(0, 1)$ .

(b) If  $\sigma = 1$  and  $\mu > 0$ , show that:  $\sqrt{n}(\ln \bar{X}_n - \ln \mu) \xrightarrow{d} V \sim N(0, \frac{1}{\mu^2})$ ;

(c) Show that  $\frac{n^\delta(\bar{X}_n - \mu)}{\sigma} \xrightarrow{P} 0$ , for any  $\delta < 0.5$ .

(d) Find the asymptotic distributions of: (i)  $\sqrt{n}(\bar{X}_n^2 - \mu^2)$ ; (ii)  $n(\bar{X}_n - \mu)^2$  and (iii)  $\sqrt{n}(\bar{X}_n - \mu)^2$ .

17. Let  $X_1, X_2, \dots$  be i.i.d. r.v.s having  $\text{Exp}(\theta)$  ( $\theta > 0$ ) distribution and let  $\bar{X}_n = \frac{\sum_{i=1}^n X_i}{n}$ ,  $n = 1, 2, \dots$ . Show that:  $\sqrt{n}(\frac{1}{\bar{X}_n} - \frac{1}{\theta}) \xrightarrow{d} N(0, \frac{1}{\theta^2})$ , as  $n \rightarrow \infty$ .

18. Let  $X_1, X_2, \dots$  be a sequence of i.i.d.  $U(0, 1)$  r.v.s. For the sequence of geometric means  $G_n = (\prod_{i=1}^n X_i)^{\frac{1}{n}}$ ,  $n = 1, 2, \dots$ , show that  $\sqrt{n}(G_n - \frac{1}{e}) \xrightarrow{d} N(0, \sigma^2)$ , for some  $\sigma^2 > 0$ . Find  $\sigma^2$ .

19. Let  $(X_1, Y_1), (X_2, Y_2), \dots$  be a sequence of independent bivariate random vectors having the same joint p.d.f. Let  $E(X_1) = \mu$ ,  $E(Y_1) = \nu$ ,  $\text{Var}(X_1) = \sigma^2$ ,  $\text{Var}(Y_1) = \tau^2$  and  $\text{Corr}(X_1, Y_1) = \rho$ . Let  $Q_n = \frac{\sum_{i=1}^n (X_i - \bar{X}_n)(Y_i - \bar{Y}_n)}{n-1}$ ,  $S_n^2 = \frac{\sum_{i=1}^n (X_i - \bar{X}_n)^2}{n-1}$ ,  $T_n^2 = \frac{\sum_{i=1}^n (Y_i - \bar{Y}_n)^2}{n-1}$  and  $R_n = \frac{Q_n}{S_n T_n}$ .

(a) Show that  $Q_n \xrightarrow{P} \rho\tau\sigma$  and  $R_n \xrightarrow{P} \rho$ .

(b) Let  $\delta = \frac{E((X_1 - \mu)^2(Y_1 - \nu)^2)}{\sigma^2\tau^2}$ . Show that  $\sqrt{n}(Q_n - \rho\sigma\tau) \xrightarrow{d} N(0, (\delta - \rho^2)\sigma^2\tau^2)$ .