

M/G/1 Queue with Vacations

Copyright 2002, Sanjay K. Bose

1

Vacation: After a *busy period*, the server goes on *vacation* of random length. It examines the queue once again when it returns from the vacation

Multiple Vacations (possibly)

If system still empty when the server returns from a vacation, it goes for another vacation. This continues until it finds system non-empty on return from vacation; it then resumes service normally

Single Vacation (per idle)

After a busy period ends, server goes on only one vacation. If system is still empty when in returns, it stays and waits for a job to arrive.

Other models are also possible, i.e. server goes on (possibly multiple) vacations following the busy period until there are K waiting jobs

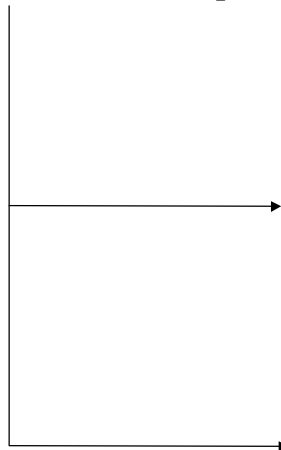
Copyright 2002, Sanjay K. Bose

2

Analysis of M/G/1 Queue with (Multiple) Vacations

Vacation Interval
random (i.i.d) and
independent of
service times with

$$\left\{ \begin{array}{l} \text{moments } \bar{V}, \bar{V}^2 \\ \text{pdf } f_V(t) \\ \text{cdf } F_V(t) \\ \text{L.T. } L_V(s) \end{array} \right.$$



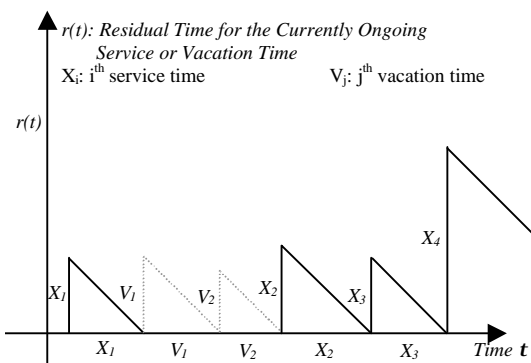
Residual Life
Approach

Imbedded
Markov Chain
Approach

Copyright 2002, Sanjay K. Bose

3

Analysis using the Residual Life based Approach



Residual Time $r(t)$ - service time or vacation - for a
M/G/1 Queue with (multiple) Vacations

Copyright 2002, Sanjay K. Bose

4

$$\text{Time Average of } r(t) \text{ over } (0, t) = \frac{1}{t} \int_0^t r(x) dx = \frac{1}{t} \sum_{i=1}^{M(t)} \frac{1}{2} X_i^2 + \frac{1}{t} \sum_{j=1}^{L(t)} \frac{1}{2} V_j^2$$

$$\text{where } \begin{cases} M(t) = \text{Number of arrivals in the interval } (0, t) \\ L(t) = \text{Number of vacation intervals in the interval } (0, t) \end{cases}$$

$$\text{For } t \rightarrow \infty \begin{cases} R = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t r(x) dx & \lim_{t \rightarrow \infty} \frac{1}{M(t)} \sum_{i=1}^{M(t)} X_i^2 = \overline{X^2} \\ \lim_{t \rightarrow \infty} \frac{M(t)}{t} = \mathbf{I} & \lim_{t \rightarrow \infty} \frac{1}{L(t)} \sum_{j=1}^{L(t)} V_j^2 = \overline{V^2} \\ \lim_{t \rightarrow \infty} \frac{t(1-r)}{L(t)} = \overline{V} \end{cases}$$

Copyright 2002, Sanjay K. Bose

5

As for the basic M/G/1 queue considered earlier, this leads to

$$\text{Mean Residual Time (service or vacation)} \quad R = \frac{1}{2} \mathbf{I} \overline{X^2} + \frac{1}{2} (1-r) \frac{\overline{V^2}}{\overline{V}} \quad (4.2)$$

$$\text{Writing} \quad W_q = N_q \overline{X} + R = \mathbf{I} W_q \overline{X} + R$$

$$\text{gives} \quad W_q = \frac{\mathbf{I} \overline{X^2}}{2(1-r)} + \frac{\overline{V^2}}{2\overline{V}} \quad \mathbf{r} = \mathbf{I} \overline{X} \quad (4.3)$$

as the mean waiting time in queue seen by an arriving customer

Knowing W_q , the other parameters N_q , N and W may be found

Copyright 2002, Sanjay K. Bose

6

Analysis using the Imbedded Markov Chain Approach

As for the basic M/G/1 queue, imbed Markov Chain of system states (denoting the number in the system) at the time instants t_i $i=1, 2, 3, \dots$ when the i^{th} customer departs from the system

n_i = Number of jobs left behind in the system by the i^{th} departure

a_i = Number of job arrivals during the i^{th} service time

$$A(z) = L_B(\mathbf{I} - \mathbf{I}z)$$

with $A(1) = 1$ $A'(1) = \mathbf{r} = \mathbf{I}\bar{X}$ $A''(1) = \mathbf{I}^2 \bar{X}^2$

Copyright 2002, Sanjay K. Bose

7

j = Number of jobs waiting for service when a busy period begins, $j \geq 1$

$f_j = P\{j \text{ customers starting the busy period}\} \quad j=1,2,\dots,\infty$

Generating Function for j
(see Sec. 4.1.1 for details)

$$\left\{ \begin{aligned} F(z) &= \sum_{j=1}^{\infty} f_j z^j = E\{z^j\} \\ &= \frac{L_V(\mathbf{I} - \mathbf{I}z) - L_V(\mathbf{I})}{1 - L_V(\mathbf{I})} \end{aligned} \right.$$

with $F(1) = 1$ $F'(1) = \frac{\mathbf{I}\bar{V}}{1 - L_V(\mathbf{I})}$ $F''(1) = \frac{\mathbf{I}^2 \bar{V}^2}{1 - L_V(\mathbf{I})}$

Copyright 2002, Sanjay K. Bose

8

Relating the state at the i^{th} and $(i+1)^{\text{th}}$ instants, we get

$$\begin{aligned} n_{i+1} &= a_{i+1} + j - 1 && \text{for } n_i = 0 \\ &= n_i + a_{i+1} - 1 && \text{for } n_i \geq 1 \end{aligned} \quad (4.5)$$

or

$$n_{i+1} = n_i + a_{i+1} - 1 + j[1 - U(n_i)] \quad (4.4)$$

$$P(z) = E\{z^{n+a-1+j[1-U(n)]}\} = E\{z^a\}E\{z^{n-1+j[1-U(n)]}\}$$

$$P(z) = A(z)E\{p_0 z^{j-1} + \sum_{n=1}^{\infty} z^{n-1} p_n\}$$

$$\Rightarrow P(z) = p_0 A(z) \frac{1 - F(z)}{A(z) - z}$$

Copyright 2002, Sanjay K. Bose

9

Evaluating $P(z) = p_0 A(z) \frac{1 - F(z)}{A(z) - z}$ at $z=1$, i.e. using $P(1)=1$, gives

$$p_0 = \frac{1 - r}{F'(1)} \quad (4.9)$$

and therefore

$$P(z) = (1 - r)(1 - L_v(\mathbf{I} - \mathbf{I}z)) \left(\frac{L_B(\mathbf{I} - \mathbf{I}z)}{z - L_B(\mathbf{I} - \mathbf{I}z)} \right) \quad (4.10)$$

Note that though $P(z)$ was derived for the customer departure instants, it will also hold for the arrival instants and at an arbitrary time instant under equilibrium conditions.

Copyright 2002, Sanjay K. Bose

10

- From $P(z)$, we can find the system state distribution either by inverting the generating function $P(z)$ or by expanding it in powers of z

- The moments of the number in the system may be found directly using the moment generating properties of the generating function $P(z)$.

- Specifically, we get $N = P'(1) = I\bar{X} + \frac{I^2 \bar{X}^2}{2(1-I\bar{X})} + \frac{IV^2}{2\bar{V}}$

- Knowing N , we can obtain W , W_q and N_q following our usual approach

For example
$$W_q = \frac{I\bar{X}^2}{2(1-r)} + \frac{\bar{V}^2}{2\bar{V}} \quad (4.11)$$

M/G/1 Queue with only one Vacation after Idle (Section 4.2)

$$P(z) = p_0 \left(\frac{L_B(I - Iz)}{z - L_B(I - Iz)} \right) \left(L_V(I - Iz) - (1 - z)L_V(I) - 1 \right)$$

with
$$p_0 = \frac{1 - I\bar{X}}{I\bar{V} + L_V(I)}$$

Using Imbedded Markov Chain

Directly using Residual Life Approach

$$W_q = \frac{I\bar{X}^2}{2(1-I\bar{X})} + \frac{\bar{V}^2}{2\left(\bar{V} + \frac{1}{I}L_V(I)\right)}$$

M/G/1 Queue with Exceptional First Service (Section 4.3)

In this queue, the first customer starting a busy period requires a service time with a different distribution, i.e. $b^*(t)$ and $L_{B^*}(s)$ with moments \bar{X}^* and \bar{X}^{*2}

An imbedded Markov Chain analysis will give

$$P(z) = \frac{p_0 [L_B(\mathbf{I} - \mathbf{I}z) - zL_{B^*}(\mathbf{I} - \mathbf{I}z)]}{L_B(\mathbf{I} - \mathbf{I}z) - z} \quad (4.19)$$

with
$$p_0 = \frac{1 - \mathbf{I}\bar{X}}{1 - \mathbf{I}\bar{X} + \mathbf{I}\bar{X}^*} \quad (4.18)$$

The delay distribution for the FCFS case, may be found using

$$P(z) = L_T(\mathbf{I} - \mathbf{I}z).$$

This may then be used to find W and W_q

Alternatively, these may be found using a *Residual Life Approach*

$$W = \frac{\bar{X}^*}{1 - \mathbf{I}\bar{X} + \mathbf{I}\bar{X}^*} + \frac{\mathbf{I}\bar{X}^2}{2(1 - \mathbf{I}\bar{X})} + \frac{\mathbf{I}(\bar{X}^{*2} - \bar{X}^2)}{2(1 - \mathbf{I}\bar{X} + \mathbf{I}\bar{X}^*)} \quad (4.21)$$

$$W_q = \frac{\mathbf{I}\bar{X}^2}{2(1 - \mathbf{I}\bar{X})} + \frac{\mathbf{I}(\bar{X}^{*2} - \bar{X}^2)}{2(1 - \mathbf{I}\bar{X} + \mathbf{I}\bar{X}^*)} \quad (4.22)$$