



# ANALYSIS OF PARTIALLY COMPLETE TIME AND TYPE OF FAILURE DATA

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# ORGANIZATIONS

- DATA DESCRIPTION
- ASSUMPTIONS
- NON-PARAMETRIC METHODS
- PARAMETRIC METHODS
- BAYESIAN TECHNIQUES
- CONCLUSIONS



## WE HAVE THE FOLLOWING OBSERVATIONS

$T = \text{FAILURE TIME} \in (0, \infty)$

$\delta = \text{CAUSE OF FAILURE} \in \{1, \dots, K\}$ .

- TYPE: 1  $\{T = t, \delta = j\}$
- TYPE: 2  $\{T > t\}$
- TYPE: 3  $\{T > t, \delta = j\}$
- TYPE: 4  $\{T = t\}$ .

IF THE DATA CONSIST OF THESE TYPES OF OBSERVATIONS, HOW TO ANALYZE THE DATA?

IT INVOLVES BY SPECIFYING A DISTRIBUTION OF  $(T, \delta)$ .



## DIFFERENT DATA SETS

### DATA SET 1: BRAIN CANCER DATA

172 PATIENTS WERE GIVEN THE RADIATION THERAPY HAVING A PARTICULAR TYPE OF BRAIN CANCER. DATA WERE RECORDED FROM THE DAY OF THE THERAPY TILL DEATH/WITHDRAWAL. ALMOST 50% HAVE A CENSORED VALUE OF  $T$ . OVER 65% HAVE AN UNKNOWN VALUE OF  $\delta$ . NO TYPE-4 DATA.

### DATA SET 1: PROSTATE CANCER DATA

79 MALE PATIENTS WERE GIVEN THE THERAPY HAVING THE STAGE 4 PROSTATE CANCER. DATA WERE RECORDED FROM THE DAY OF THE THERAPY TILL DEATH/WITHDRAWAL. ALMOST 35% HAVE A CENSORED VALUE OF  $T$ . OVER 50% HAVE AN UNKNOWN VALUE OF  $\delta$ . 16% BELONG TO TYPE-4.



## NON-PARAMETRIC APPROACH

LET US DEFINE:

$$S(t) = P(T > t) \quad \delta_j(t) = P(\delta = j | T = t).$$

LIKELIHOOD CONTRIBUTION:

- TYPE 1:  $\delta_j(t)dS(t)$
- TYPE 2:  $S(t)$
- TYPE 3:  $\int_t^\infty \delta_j(u)dS(u)$
- TYPE 4:  $dS(t)$

$$dS(t) = \begin{cases} -S'(t)dt & \text{IF } T \text{ IS ABSOLUTELY CONT.} \\ S(t-) - S(t) & \text{IF } T \text{ IS DISCRETE} \end{cases}$$

ASSUMPTION: MISSING  $\delta$  HAPPENS AT RANDOM



## NON-PARAMETRIC MLEs

SUPPOSE  $t_1, \dots, t_K$  ARE THE OBSERVED TIME POINTS WHERE SOME EVENTS HAVE TAKEN PLACE. AT EACH TIME POINT  $t_k$  DIVIDE THE NUMBER OF OBSERVATIONS INTO FOUR GROUPS.

$$S(t_m) = \prod_{l=1}^m p_l, \quad dS(t_m) = (1 - p_m) \prod_{l=1}^{m-1} p_l$$

$$p_l = \frac{S(t_l)}{S(t_{l-1})}, \quad \delta_{jm} = \delta_j(t_m)$$

$$\delta_{j,K+1} = P(\delta = j | T > t_K)$$

IT IS POSSIBLE TO OBTAIN THE MLEs USING THE EM ALGORITHM



## PARAMETRIC METHOD ( $T$ IS ASSUMED TO BE CONTINUOUS)

### LATENT FAILURE TYPE MODEL (COX)

$T_j$  = THE RANDOM VARIABLE IS THE TIME OF FAILURE  
IF THE INDIVIDUAL FAILS DUE TO CAUSE  $j$

WE DO NOT OBSERVE  $T_1, \dots, T_K$ .

WE OBSERVE:

$$T = \min\{T_1, \dots, T_K\}$$

$$\delta = \{j; T_j < T_l, l = 1, \dots, K, l \neq j\}$$

IT IS ASSUMED THAT  $T_1, \dots, T_K$  ARE INDEPENDENT. IT  
IS AN ASSUMPTION AND IT CAN NOT BE VERIFIED FROM  
THE OBSERVED DATA

FOR PARAMETRIC STUDIES,  $T_i$ 'S ARE ASSUMED TO  
HAVE DIFFERENT PARAMETRIC FORMS



# CAUSE SPECIFIC HAZARD FUNCTION APPROACH

OVERALL HAZARD FUNCTION

$$\lambda(t) = \lim_{\Delta t \rightarrow 0} \frac{P(t \leq T < t + \Delta t | T \geq t)}{\Delta t}$$

CAUSE SPECIFIC HAZARD FUNCTION

$$\lambda_j(t) = \lim_{\Delta t \rightarrow 0} \frac{P(t \leq T < t + \Delta t, \delta = j | T \geq t)}{\Delta t}$$

$$\lambda(t) = \sum_{j=1}^K \lambda_j(t) \quad S(t) = e^{-\int_0^t \lambda(u) du}$$

THE PROBABILITY DENSITY FUNCTION DUE TO CAUSE  $j$  CAN BE WRITTEN AS

$$f_j(t) = \lambda_j(t)S(t)$$

DIFFERENT PARAMETRIC FORMS OF  $\lambda_j$ 'S ARE ASSUMED. IT IS OBSERVED THAT IN CASE OF EXPONENTIAL AND WEIBULL MODELS THE TWO FORMULATIONS ARE EQUIVALENT.



# LATENT FAILURE TIME MODEL

EXPONENTIAL MODEL:

SUPPOSE  $T_j$ 'S ARE INDEPENDENT EXPONENTIAL RANDOM VARIABLES WITH MEAN  $\frac{1}{\lambda_j} = \theta_j$

IT IS POSSIBLE TO OBTAIN THE MLES AND UMVUES OF  $\lambda_j$ 'S

THE MLE OF  $\theta_j$  DOES NOT EXIST. WE PROPOSE TO USE THE CONDITIONAL MLES. THE CONDITIONAL PROBABILITY DENSITY FUNCTIONS OF  $\hat{\theta}_j$  IS A MIXTURE OF GAMMA DENSITY FUNCTIONS. IT CAN BE USED TO CONSTRUCT THE CONFIDENCE INTERVAL ALSO.



## WEIBULL FAILURE DISTRIBUTION

THE DISTRIBUTION FUNCTION OF  $T_j$ 'S HAVE THE FOLLOWING FORM:

$$F_j(t) = 1 - e^{-\lambda_j t^\alpha}$$

THE MLES OF  $\alpha$  AND  $\lambda_1, \dots, \lambda_K$  DO NOT HAVE EXPLICIT FORMS. EM ALGORITHM CAN BE USED TO COMPUTE THE MLES BUT IT TAKES A LONG TIME TO CONVERGE. A FIXED POINT TYPE ALGORITHM HAS BEEN PROPOSED AND THE CONVERGENCE IS QUITE FAST.



# PARAMETRIC AND NON-PARAMETRIC ESTIMATIONS

## DATA SET 1

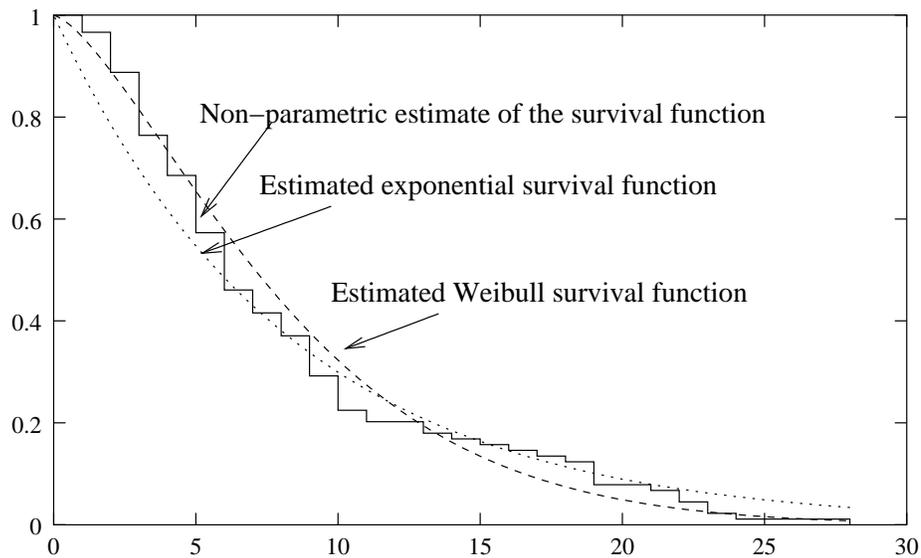


Figure 1: Empirical survival function and the estimated survival functions using exponential and Weibull models.



## DATA SET 2

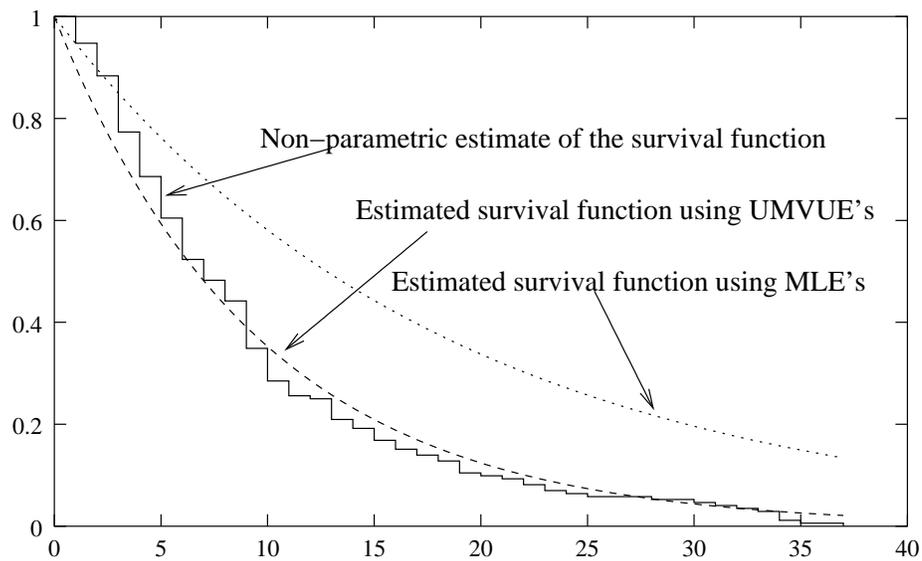


Figure 2: Empirical survival function and estimated survival functions using MLE's and UMVUE's. Here lifetime distributions are exponential.



## DATA SET 2

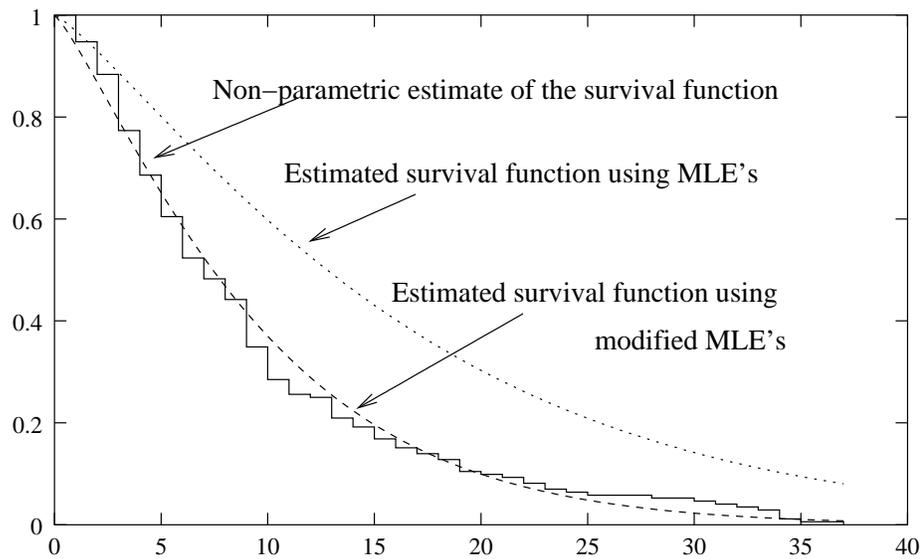


Figure 3: Empirical survival function and the estimated survival functions using MLE's and modified MLE's. Here the lifetime distributions are Weibull.



# BAYESIAN ANALYSIS

## EXPONENTIAL MODEL

$\lambda_j$ S HAVE GAMMA INDEPENDENT PRIORS

TWO CASES DEPENDING ON ( $\#$  TYPE 4 -  $\#$  TYPE 3)  
OBSERVATIONS

- IF ( $\#$  TYPE 4 -  $\#$  TYPE 3)  $\geq 0$  THEN EXPLICIT SOLUTIONS
- IF ( $\#$  TYPE 4 -  $\#$  TYPE 3)  $< 0$  THEN NO EXPLICIT SOLUTION
- IF ( $\#$  TYPE 4 -  $\#$  TYPE 3) = 0 THEN VERY CONVENIENT FORM AND EXPLICIT CREDIBLE INTERVALS
- IF ( $\#$  TYPE 4 -  $\#$  TYPE 3)  $> 0$  THEN IT CAN BE OBTAINED AS MIXTURES AND CONSERVATIVE CREDIBLE INTERVALS CAN BE CONSTRUCTED



## WEIBULL MODEL

$\lambda_j$ S HAVE INDEPENDENT GAMMA PRIORS AND WE DO NOT NEED TO ASSUME ANY SPECIFIC FORM OF THE PRIOR ON  $\alpha$ . IT IS ASSUMED THAT IT HAS THE SUPPORT ON  $(0, \infty)$  AND IT HAS A LOG-CONCAVE DENSITY FUNCTION.

THE POSTERIOR DENSITY FUNCTIONS OF  $\lambda_j$  GIVEN  $\alpha$  CAN BE OBTAINED EXPLICITLY.

THE POSTERIOR DENSITY FUNCTION OF  $\alpha$  IS LOG-CONCAVE