

4. IMPRECISE PROBABILITY AND UNCERTAIN EVIDENCE

1. General framework and inference
with imprecise probability
2. Random sets and Belief functions
3. Merging uncertain information

A GENERAL SETTING FOR REPRESENTING GRADED CERTAINTY AND PLAUSIBILITY

- 2 adjoint set-functions Pl and Cr.
- **Conventions :**
 - $Pl(A) = 0$ "impossible" ;
 - $Cr(A) = 1$ "certain"
 - $Pl(A) = 1 ; Cr(A) = 0$ "ignorance" (**no information**)
 - $Cr(A) \leq Pl(A)$ "certain implies plausible"
 - $Pl(A) = 1 - Cr(A^c)$ duality certain/plausible

Imprecise probability theory

- A state of information is represented by a family \mathcal{P} of probability distributions over a set X .
- To each event A is attached a probability interval $[P_*(A), P^*(A)]$ such that
 - $P_*(A) = \inf\{P(A), P \in \mathcal{P}\}$
 - $P^*(A) = \sup\{P(A), P \in \mathcal{P}\} = 1 - P_*(A^c)$

Subjectivist view (Peter Walley)

- $P_{\text{low}}(A)$ is the highest acceptable price for buying a bet on event A winning 1 euro if A occurs
- $P^{\text{high}}(A) = 1 - P_{\text{low}}(A^c)$ is the least acceptable price for selling this bet.
- **Coherence** condition

$$P_*(A) = \inf\{P(A), P \geq P_{\text{low}}\} = P_{\text{low}}(A)$$

Imprecise probability theory

- The most general numerical approach to uncertainty.
 - **Attention !** If $\text{Cr}(A)$, $\text{Pl}(A)$ derive from a family \mathcal{P}
 $\mathcal{P} \neq \{ P, P(A) \in [\text{Cr}(A), \text{Pl}(A)] \text{ for all } A \}$;
 - Only $\mathcal{P} \subset \{ P, P(A) \in [\text{Cr}(A), \text{Pl}(A)] \text{ for all } A \}$ holds
- Equality if $\text{Cr}(A) + \text{Cr}(B) \leq \text{Cr}(A \cup B) + \text{Cr}(A \cap B)$
for all A, B (super-additivity of convex capacities)

PROBABILISTIC LOGIC (de Finetti)

- A probabilistic knowledge base is a set of weighted propositions $\mathcal{B} = \{(p_i, \alpha_i), i = 1, \dots, \alpha_i \in [0, 1]\}$
- (p_i, α_i) means that $P(A_i) = \alpha_i$ where $A_i = [p_i]$
= A constraint on an unknown probability distribution on interpretations.
- \mathcal{B} is not enough to isolate a single probability distribution :
it defines a probability family on interpretations
 - $\mathcal{P} = \{P: P(A_i) = \alpha_i, i = 1, \dots\}$
- ***Inference problem:*** *What can be said about the probability another proposition p ???*
- *All that is known is $P(p) \in [\alpha, \beta]$, generally.*

Probabilistic extension of propositional logic

- *Finding the most narrow interval $[\alpha, \beta]$ such that $\mathcal{B} \models P(p) \in [\alpha, \beta]$ is a **non-straightforward** extension of classical inference.*
 - *More generally \mathcal{B} may contain information of the form $P(A_i) \in [\alpha_i, \beta_i]$*
 - *It cannot distinguish between contingent evidence and generic knowledge*
 - *It cannot model conditional probability*
 - *Information pieces of the form $(A_i \alpha_i)$ and $(A_i \Rightarrow B_i \beta_i)$, with material implication \Rightarrow are not independent*

CONDITIONAL PROBABILISTIC LOGIC

- A conditional probabilistic knowledge base is a weighted set of rules $\Delta = \{(p_i \rightarrow q_i, \alpha_i), i = 1, \dots\}$ where α_i lies in $[0, 1]$.
 - **It represents Generic knowledge**
 - $(p_i \rightarrow q_i, \alpha_i) \in \Delta$ means that $P(q_i | p_i) = \alpha_i$
 - A probability family $\mathcal{P}_\Delta = \{P, P(q_i | p_i) = \alpha_i, i = 1, \dots\}$ is the semantic counterpart of Δ .
- *Mathematically it is more complex than Bayesian nets but it has the power of probability theory and leaves room to incompleteness.*

Inference with a probabilistic conditional base

- *Suppose a set of probabilistic conditionals Δ and C a propositional base. Let $\mathcal{C} = [C]$ be the set of models.*
- *Two types of processing :*
 - 1. **Querying** : \mathcal{C} is a set of singular facts : compute the degree of belief of A in context \mathcal{C} as*
$$\text{Cr}(A | \mathcal{C}) = \text{Inf}\{P(A | C), P \in \mathcal{P}_\Delta, P(C) > 0 \}.$$
 - 2. **Revision** : \mathcal{C} is a set of universal truths;*

Add $P(C) = 1$ to the set of conditionals \mathcal{P}_Δ .

Now we must compute $\text{Cr}(A | \mathcal{C}) = \text{Inf}\{P(B) \mid P \in \mathcal{P}_\Delta, P(C) = 1 \}$

If $P(C) = 1$ is incompatible with \mathcal{P}_Δ , consider

$$\text{Cr}(A | \mathcal{C}) = \text{Inf}\{P(B | C) \mid P \in \mathcal{P}_\Delta, P(C) \text{ maximal} \}$$

Example : $A \xleftrightarrow{\quad} B \longrightarrow C$

- \mathcal{P} is the set of probabilities such that
 - $P(B|A) \geq \alpha$ *Most A are B*
 - $P(C|B) \geq \beta$ *Most B are C*
 - $P(A|B) \geq \gamma$ *Most B are A*
- **Querying on context A** : Find the most narrow interval for $P(C|A)$ (*Linear programming*): we find
$$P(C|A) \geq \beta \cdot \max(0, 1 - (1 - \gamma)/\alpha)$$
 - *Note* : if $\gamma = 0$, $P(C|A)$ is unknown even if $\alpha = 1$.
- **Revision** : Suppose $P(A) = 1$, then $P(C|A) \geq \beta \cdot \gamma$
 - *Note* : $\gamma > \max(0, 1 - (1 - \gamma)/\alpha)$
- **Revision improves generic knowledge, querying does not.**

Imprecise probability inference extend preferential (cautious) inference

- *Infinitesimal probabilistic inference*: From set of rules Δ of the form $P(q_i | p_i) \geq 1 - \varepsilon$, prove that $P(q | p) \geq 1 - O(\varepsilon)$ from Δ where ε is an infinitesimal number.
- This is equivalent to proving $\Delta^* \models_{\forall} p \rightarrow q$ using cautious possibilistic inference (or Lehman's system P) where $\Delta^* = \{p_i \rightarrow q_i\}$ is the set of conditionals present in Δ (Adams' logic of conditionals, 1975).
- It is equivalent to the logic of conditional events
- It is equivalent to inference from set of rules Δ of the form $P(q_i | p_i) = 1$ under De Finetti's coherence approach (T. Lukasiewicz, A. Gilio).

Random sets and evidence theory

- A family \mathcal{F} of « focal » (disjunctive) non-empty sets representing
 - Statistics under incomplete observations (generic)
 - Unreliable testimonies (singular)
 - Indirect information (induced from a probability space)
- A positive weighting of focal sets (a random set) :
$$\sum_{E \in \mathcal{F}} m(E) = 1 \quad (\textit{mass function})$$
- It is a randomized incomplete information

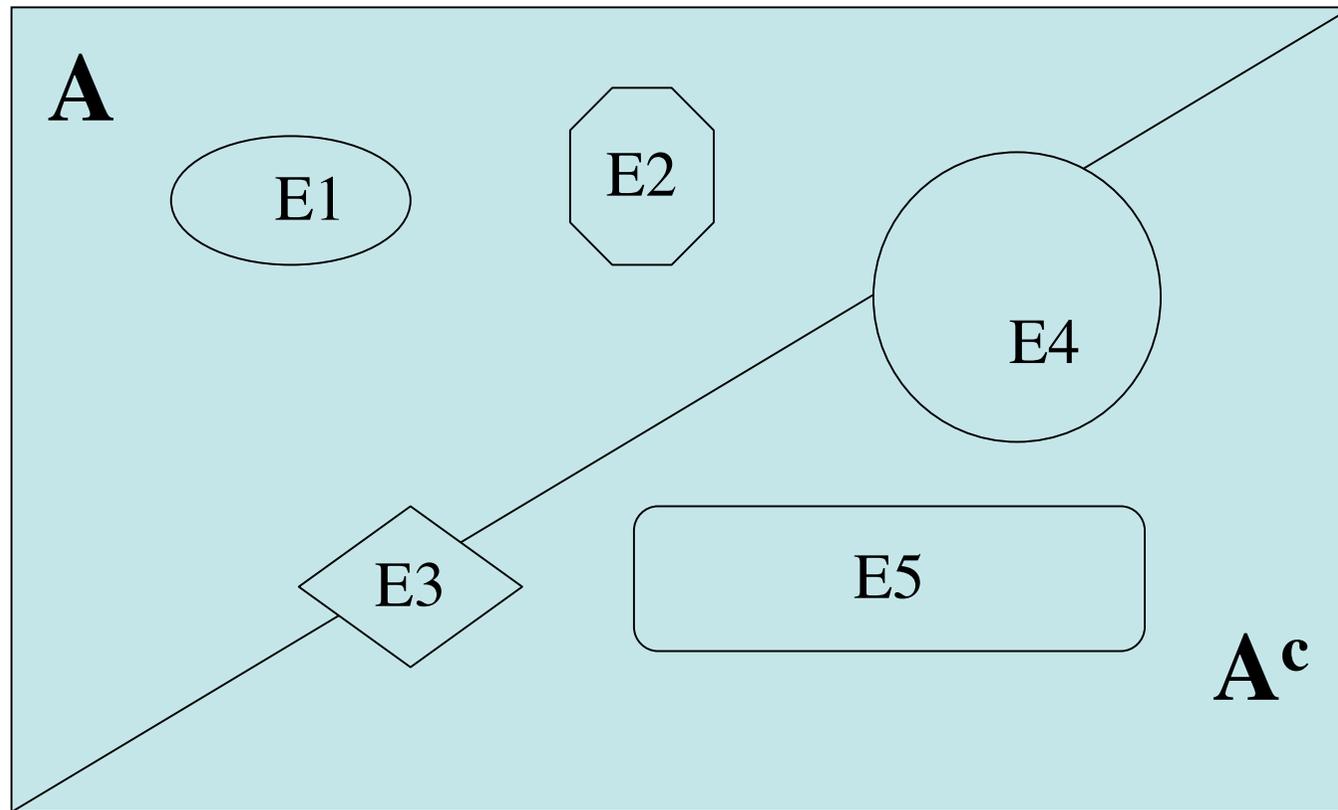
Theory of evidence

- $m(E)$ = probability that the most precise description of the available information is of the form " $x \in E$ "
 - = probability (only knowing " $x \in E$ " and nothing else)*
 - It is the portion of probability mass hanging over elements of E without being allocated.
- **DO NOT MIX UP $m(E)$ and $P(E)$**
- In the view of Shafer (1976) and Smets the mass assignment m represents uncertain singular evidence on the solution of a problem

Theory of evidence

- **degree of certainty (belief) :**
 - (Function Cr denoted Bel by Shafer)
 - $\text{Bel}(A) = \sum_{E_i \subseteq A, E_i \neq \emptyset} m(E_i)$
 - total mass of information implying the occurrence of A
 - (*probability of provability*)
- **degree of plausibility :**
 - $\text{Pl}(A) = \sum_{E_i \cap A \neq \emptyset} m(E_i) = 1 - \text{Bel}(A^c) \geq \text{Bel}(A)$
 - total mass of information consistent with A
 - (*probability of consistency*)

Example : $\text{Bel}(A) = m(E1) + m(E2)$
 $\text{Pl}(A) = m(E1) + m(E2) + m(E3) + m(E4)$
 $= 1 - m(E5) = 1 - \text{Bel}(A^c)$



PARTICULAR CASES

- INCOMPLETE INFORMATION:

$$m(E) = 1, m(A) = 0, A \neq E$$

- *TOTAL IGNORANCE* : $m(S) = 1$:

– *For all $A \neq S, \emptyset, Bel(A) = 0, Pl(A) = 1$*

- PROBABILITY : if $\forall i, E_i = \text{singleton } \{s_i\}$ (hence disjoint focal sets)

– Then, *for all $A, Bel(A) = Pl(A) = P(A)$*

– *Hence precise + scattered information*

- POSSIBILITY THEORY : the opposite case (ZADEH)

$E_1 \subseteq E_2 \subseteq E_3 \dots \subseteq E_n$: imprecise and coherent information

– iff $Pl(A \cup B) = \max(Pl(A), Pl(B))$, possibility measure

– iff $Bel(A \cap B) = \min(Bel(A), Bel(B))$, necessity measure

Plausible states induced by a belief function

- A random set (\mathcal{F}, m) expressing available information on x induces a fuzzy set F of plausible values of x , summing for each value s , masses of all focal sets containing s .
- $\mu_F(s) = \text{Pl}(\{s\}) = \sum \{m(E_i), E_i \in \mathcal{F}, s \in E_i\}$
 - $\exists s, \mu_F(s) = 1$ iff $s \in E_i$ for all i
 - $\bigcap_i E_i \neq \emptyset$ (no conflict in information).
- The mass function cannot be reconstructed from function μ_F , except if
 - All E_i are disjoint : $m(E_i) = \mu_F(s)$ if $s \in E_i$
 - E_i are all nested : $\text{Pl}(\{s\})$ is a possibility distribution

Example of uncertain evidence : Unreliable testimony (SHAFER-SMETS VIEW)

- « John tells me the president is between 60 and 70 years old, but there is some chance (*subjective probability p*) he does not know and makes it up».
 - $E = [60, 70]$; $\text{Prob}(\text{Knowing } "x \in E = [60, 70]") = 1 - p$.
 - With probability p , John invents the info, so *we know nothing (Note that this is different from a lie)*.
- We get a *simple support belief function* :
 $m(E) = 1 - p$ and $m(S) = p$
- Equivalent to a possibility distribution
 - $\pi(s) = 1$ if $x \in E$ and $\pi(s) = p$ otherwise.

CONDITIONING UNCERTAIN SINGULAR EVIDENCE

- A mass function m on S , represents *uncertain evidence*
 - A new **sure** piece of evidence is viewed as a conditioning event C
1. *Mass transfer* : for all $E \in \mathcal{F}$, $m(E)$ moves to $C \cap E \subseteq C$
 - The mass function after the transfer is $m_t(B) = \sum_{E: C \cap E = B} m(E)$
 - But the mass transferred to the empty set may not be zero!
 - $m_t(\emptyset) = \text{Bel}(C^c) = \sum_{E: C \cap E = \emptyset} m(E)$ is the degree of conflict with evidence C
 2. *Normalisation* : $m_t(B)$ should be divided by
$$\text{Pl}(C) = 1 - \text{Bel}(C^c) = \sum_{E: C \cap E \neq \emptyset} m(E)$$
- *This is revision of an unreliable testimony by a sure fact*

DEMPSTER RULE OF CONDITIONING = PRIORITIZED MERGING

The conditional plausibility function $Pl(.|E)$ is

- $Pl(A|E) = \frac{Pl(A \cap E)}{Pl(E)}$; $Bel(A|E) = 1 - Pl(A^c|E)$

- E surely contains the value of the unknown quantity described by m. **So $Pl(E^c) = 0$**
 - *The new information is interpreted as asserting the impossibility of E^c : Since E^c is impossible you can change $x \in C$ into $x \in C \cap E$ and transfer the mass of focal set C to $C \cap E$.*
- *The new information improves the precision of the evidence*

EXAMPLE OF REVISION OF EVIDENCE :

The criminal case

- **Evidence 1** : three suspects : Peter Paul Mary
- **Evidence 2** : The killer was randomly selected man vs.woman by coin tossing.
 - So, $S = \{ \text{Peter, Paul, Mary} \}$
- **TBM modeling** : The masses are $m(\{ \text{Peter, Paul} \}) = 1/2$; $m(\{ \text{Mary} \}) = 1/2$
 - $\text{Bel}(\text{Paul}) = \text{Bel}(\text{Peter}) = 0$. $\text{Pl}(\text{Paul}) = \text{Pl}(\text{Peter}) = 1/2$
 - $\text{Bel}(\text{Mary}) = \text{Pl}(\text{Mary}) = 1/2$
- **Bayesian Modeling**: A prior probability
 - $P(\text{Paul}) = P(\text{Peter}) = 1/4$; $P(\text{Mary}) = 1/2$

- **Evidence 3** : Peter was seen elsewhere at the time of the killing.
- **TBM**: So $P(\text{Peter}) = 0$.
 - $m(\{\text{Peter}, \text{Paul}\}) = 1/2$; $m_t(\{\text{Paul}\}) = 1/2$
 - *A uniform probability on {Paul, Mary} results.*
- **Bayesian Modeling**:
 - $P(\text{Paul} \mid \text{not Peter}) = 1/3$; $P(\text{Mary} \mid \text{not Peter}) = 2/3$.
 - A very debatable result that depends on where the story starts.
- *Starting with i males and j females:*
 - $P(\text{Paul} \mid \text{Paul OR Mary}) = j/(i + j)$;
 - $P(\text{Mary} \mid \text{Paul OR Mary}) = i/(i + j)$

THE IMPRECISE PROBABILITY VIEW (Dempster, 1967)

- A belief function on S is induced by a probability space (Ω, P) via a point to set-mapping $G: m(E_i) = p(w_i)$ if $G(w_i) = E_i$.
- Consider a selection function $\phi: \Omega \rightarrow S$ from G . For each focal set E_i assign mass $m(E_i)$ to element $\phi(w_i) \in E_i$: We get a probability P^ϕ such that

$$P^\phi(\{s\}) = \sum \{p(w_i) , \phi(w_i) = s\}$$

- $\forall A, \text{Bel}(A) \leq P^\phi(A)$ and (equivalently) $P^\phi(A) \leq \text{Pl}(A)$
 - $\mathcal{P} = \{P, \forall A, \text{Bel}(A) \leq P(A)\}$ = convex hull of probabilities P^ϕ .
 - $\text{Cr} = P_*$: lower probability; $\text{Pl} = P^*$: upper probability
- $P \in \mathcal{P}$ is of the form : $P(A) = \sum P_i(A|E_i) \cdot m(E_i)$ (where P_i is any probability measure on support E_i)

Theory of evidence vs. imprecise probabilities

- The set $\mathcal{P}_{\text{bel}} = \{P \geq \text{Bel}\}$ is coherent: Bel is a special case of lower probability
- Bel is ∞ -monotone (super-additive at any order)
- The solution m to the set of equations $\forall A \subseteq X$

$$g(A) = \sum_{E_i \subseteq A, E_i \neq \emptyset} m(E_i)$$

is unique (Moebius transform)

- **It is positive iff g is a belief function**

Indirect information: The unreliable watch

- **States of the watch:** $\Omega = \{OK, KO\}$;
- **Space of interest for the agent :** $S = \{\text{day hours}\}$
- **Available information:**
 1. *on the watch state:* $p = \text{Prob}(KO)$ (small)
 2. *Logical relation between Ω and S :*
 - $G(OK) = \{\text{Actual hour up to 2 mn}\} = H \subseteq S$.
 - $G(KO) = S$ (ignorance) (*broken watches may give the right time*)
- There is a probability p that agent ignores the right time : $m(H) = 1 - p$ and $m(S) = p$

Example of generic belief function: imprecise observations in an opinion poll

- **Question** : who is your preferred candidate
in $C = \{a, b, c, d, e, f\}$???
 - **To a population** $\Omega = \{1, \dots, i, \dots, n\}$ of n persons.
 - **Imprecise responses** $\mathbf{r} = \langle x(i) \in E_i \rangle$ **are allowed**
 - No opinion ($r = C$) ; « left wing » $r = \{a, b, c\}$;
 - « right wing » $r = \{d, e, f\}$; a moderate candidate : $r = \{c, d\}$
- **Definition of mass function**:
 - $m(E) = \text{card}(\{i, E_i = E\})/n$
 - = Proportion of imprecise responses $\langle x(i) \in E \rangle$

- *The probability that a candidate in subset $A \subseteq C$ is elected is imprecise :*

$$\text{Bel}(A) \leq P(A) \leq \text{Pl}(A)$$

- **There is a fuzzy set F of potential winners:**

$$\mu_F(x) = \sum_{x \in E} m(E) = \text{Pl}(\{x\})$$

- $\mu_F(x)$ is an upper bound of the probability that x is elected. It gathers responses of those who *did not give up voting* for x
- $\text{Bel}(\{x\})$ gathers responses of those who claim they will vote for x and no one else.

CONDITIONING IMPRECISE PROBABILISTIC INFORMATION

- A disjunctive random set (\mathcal{F}, m) representing background knowledge is equivalent to a set of probabilities $\mathcal{P} = \{P : \forall A, Pl(A) \geq P(A) \geq Bel(A)\}$ (*NOT conversely*).
- Querying this information based on evidence E comes down to performing a sensitivity analysis on the conditional probability $P(\cdot|E)$
 - $Bel_E(A) = \inf \{P(A|E) : P \in \mathcal{P}, P(A) > 0\}$
 - $Pl_E(A) = \sup \{P(A|E) : P \in \mathcal{P}, P(A) > 0\}$
- **This conditioning is different from Dempster conditioning**

- **Theorem:** functions $\text{Bel}_E(A)$ and $\text{Pl}_E(A)$ are belief and plausibility functions of the form
- $\text{Bel}_E(A) = \text{Bel}_E(E \cap A) / (\text{Bel}_E(E \cap A) + \text{Pl}_E(E \cap A^c))$
- $\text{Pl}_E(A) = \text{Pl}_E(E \cap A) / (\text{Pl}_E(E \cap A) + \text{Bel}_E(E \cap A^c))$
- $\text{Bel}_E(A) = 1 - \text{Pl}_E(A^c)$
- *They are less informative than Dempster conditioning: If $E \cap C \neq \emptyset$ and $E \cap C^c \neq \emptyset$ for all $C \in \mathcal{F}$, then $m_E(E) = 1$ (total ignorance on E)*
 - **Example: If opinion poll yields :** $m(\{a, b\}) = \alpha$, $m(\{c, d\}) = 1 - \alpha$, the proportion of voters for a candidate in $E = \{b, c\}$ is unknown.
 - *However if we hear a and d resign ($\text{Pl}(\{a, d\}) = 0$) then $m(\{b\}) = \alpha$, $m(\{c\}) = 1 - \alpha$ (Dempster conditioning)*

Quantitative possibility theory

- **Membership functions of fuzzy sets**
 - Natural language descriptions pertaining to numerical universes (fuzzy numbers)
 - Results of fuzzy clustering

Semantics: metrics, proximity to prototypes

- **Imprecise probability**
 - Random experiments with imprecise outcomes
 - Special convex probability sets

Semantics: frequentist, or subjectivist (gambles)...

- **Order of magnitude of extreme probabilities** (Spohn functions with values on integers)

Quantitative possibility theory

- **Likelihood functions** $\lambda(x) = P(A| x)$ behave like possibility distributions when there is no prior on x , and $\lambda(x)$ is used as the likelihood of x .
- It holds that $\lambda(B) = P(A| B) \leq \max_{x \in B} P(A| x)$
- If $P(A| B) = \lambda(B)$ then λ should be set-monotonic:
 $\{x\} \subseteq B$ implies $\lambda(x) \leq \lambda(B)$

It implies $\lambda(B) = \max_{x \in B} \lambda(x)$

POSSIBILITY AS UPPER PROBABILITY

- Given a numerical possibility distribution π , define

$$\mathcal{P}(\pi) = \{P \mid P(A) \leq \Pi(A) \text{ for all } A\}$$

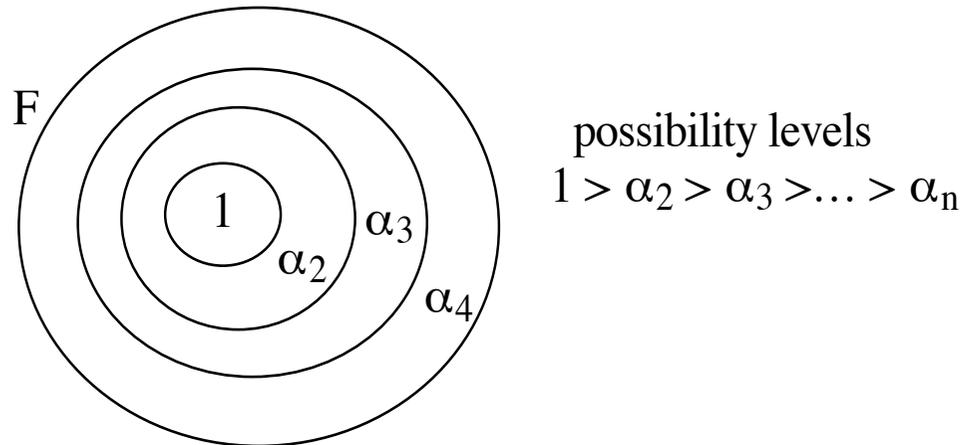
- Then, generally it holds that

$$\Pi(A) = \sup \{P(A) \mid P \in \mathcal{P}(\pi)\};$$

$$N(A) = \inf \{P(A) \mid P \in \mathcal{P}(\pi)\}$$

- So π is a faithful representation of a family of probability measures

Random set view



- Let $m_i = \alpha_i - \alpha_{i+1}$ then $m_1 + \dots + m_n = 1$,
with focal sets = cuts

A basic probability assignment (SHAFER)

- $\pi(s) = \sum_{i: s \in F_i} m_i$ (one point-coverage function) = $Pl(\{s\})$.
- *Only in the consonant case can m be recalculated from π*
- $Bel(A) = \sum_{F_i \subseteq A} m_i = N(A)$; $Pl(A) = \Pi(A)$

LANDSCAPE OF UNCERTAINTY THEORIES

BAYESIAN/STATISTICAL PROBABILITY

Randomized points

(extreme probabilities)

UPPER-LOWER PROBABILITIES

Disjunctive sets of probabilities



DEMPSTER UPPER-LOWER PROBABILITIES

SHAFER-SMETS **BELIEF FUNCTIONS**

Random disjunctive sets



Quantitative Possibility theory

Fuzzy (nested disjunctive) sets

KAPPA FUNCTIONS

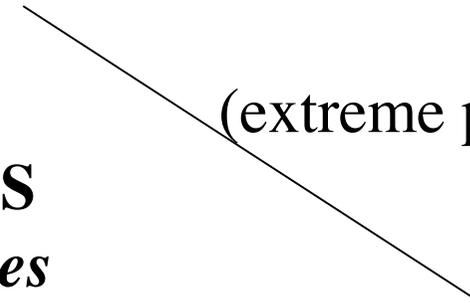
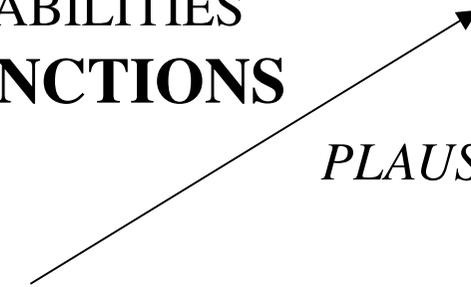
(SPOHN)

PLAUSIBILITY RANKING



Classical logic

Disjunctive sets



UNCERTAIN INFORMATION MERGING

- **Contexts :**
 - experts; sensors; images;
 - belief sets; databases; sets of propositions.
- Neither classical logic nor probability theory explain how to combine conflicting information.
- Merging beliefs differs from preference aggregation, revision.
- Theories (probability, possibility, random sets, etc...) supply connectives without explaining how to use them
- The problem is independent from the chosen representation.

WORKING ASSUMPTIONS

- Parallel information sources
- Sources are identified, heterogeneous, dependent (humans, sensors.)
- A range of problems : informing about the value of some ill-known quantity to the identification of a scenario
- Information can be poor (intervals, linguistic), incomplete, ordinal
- No prior knowledge must be available
- Reliability of sources possibly unknown, or not quantified
- Sources supposedly refer to the same problem (non-trivial issue)

BASIC MERGING MODES

source 1 : $x \in A$

$x \in ?$ *3 basic possibilities*

source 2 : $x \in B$

1. Conjunctive merging: $x \in A \cap B$

- Assumption : sources are totally reliable
- Usual in logic if no contradiction ($A \cap B \neq \emptyset$)

2. Disjunctive merging: $x \in A \cup B$

- Assumption : one of the two sources is reliable
- Imprecise but sure response : $A \cap B \neq \emptyset$ is allowed

BASIC MERGING MODES

3. Merging by counting:

build the random set : $m(A) = m(B) = 1/2$.

- $AMB(x) = Pl(x) = \sum_{x \in E} m(E) = 1$ if $x \in A \cap B$
 $= 1/2$ if $x \in (A^c \cap B) \cup (A \cap B^c)$
 $= 0$ otherwise
 - It lies between conjunctive and disjunctive (but *AMB is a fuzzy set*) : $A \cap B \subseteq AMB \subseteq A \cup B$
 - Assumption : Pieces of information stem from identical independent sources: confirmation effect.
 - Usual assumption in statistics with many sources and precise observations

Extension to n sources : conflict management with incomplete information

- A set S of n sources $i : x_i \in A_i, i = 1, \dots, n$
 - Generally inconsistent so conjunctive merging fails
 - *Significant dispersion so disjunctive merging is uninformative*
 - (there is often more than one reliable source among n)
- **Method 1** : Find maximal consistent subsets of sources $\mathcal{T}_k : \bigcap_{i \in \mathcal{T}} A_i \neq \emptyset$ but $\bigcap_{i \in \mathcal{T} \cup \{j\}} A_i = \emptyset$
 - Conjunctive merging of information in \mathcal{T}_k
 - Disjunctive merging of partial results obtained

$$X = \bigcup_k \left(\bigcap_{i \in \mathcal{T}_k} A_i \right)$$

- **Method 2** : Make an assumption on the number of reliable sources
- *Suppose k reliable sources*
- Then pick k sources at random for conjunctive merging and then disjunctively merge obtained results

$$X = \bigcup_{\mathcal{K} \subseteq S : \text{card}(\mathcal{K}) = k} \bigcap_{i \in \mathcal{K}} A_i$$

– Must choose $k \leq \max \{ \text{card}(\mathcal{K}), \bigcap_{i \in \mathcal{K}} A_i \neq \emptyset \}$

- **Method 3** : statistical : $m(A_i) = 1/n$ for all i.

$$\text{then } Pl(x) = \sum_{i=1, \dots, n} A_i(x)/n.$$

MERGING IN POSSIBILITY THEORY:

- *Fuzzy set-theoretic operations are instrumental.*
- General case :
- source 1 $\rightarrow \pi_1 = \mu_{F^1}$ source 2 $\rightarrow \pi_2 = \mu_{F^2}$
- 1. **Conjunctive merging** $F_1 \cap F_2$
 - Assumption 1 : Nothing is assumed about dependence of sources
 - **Then, Idempotence**: no accumulation effect :
- $\pi_{\cap} = \min(\pi_1, \pi_2)$ (minimum rule)
- In agreement with the logical view of information as constraints

Normalized conjunctive merging

- **Degree of conflict** : $1 - \max \pi_{\cap}$ if π_{\cap} is not normalized
 - **Renormalizing** : Assumption 2: sources are reliable even if conflict.
- Assumptions 1 and 2 : $\pi_{\cap}^* = \min(\pi_1, \pi_2) / \max \pi_{\cap}$
 - *But then Associativity is lost*
- Assumption 3: **Independent sources**: $\pi_* = \pi_1 \cdot \pi_2$
 - product instead of min.
 - **Renormalizing** : $\pi = \pi_1 \cdot \pi_2 / \max \pi_*$
 - in agreement with the Bayesian approach.
 - *Associativity is preserved*

- **Possibilistic disjunctive merging**
 - Assumption 4: one of the sources is reliable
 - $F_1 \cup F_2 : \pi_{\cup} = \max(\pi_1, \pi_2)$ (max rule)
 - **Idempotent**: sources can be redundant.
 - Adapted for inconsistent sources ($F_1 \cap F_2 = \emptyset$)

- **Statistical Merging** *vertical average*

$$\pi_+ = (\pi_1 + \pi_2) / 2$$
 - Assumption 5: Numerous identical independent sources
 - *Generally it gives a random fuzzy set.*

MERGING PROBABILITY DISTRIBUTIONS

- The basic connective is the *convex combination* :
a counting scheme
 - $P_1 \dots P_n$ probability distributions
 - Information sources with weights α_i such that $\sum \alpha_i = 1$
$$P = \sum \alpha_i P_i$$
- The only possible one with
 - $P(A) = f(P_1(A), \dots, P_n(A)) \quad \forall A \subseteq S$
 - $f(0, 0 \dots 0) = 0$; $f(1, 1 \dots 1) = 1$
 - (invariant via projections)
- Information items come from a random source ;
weights express repetition of sources: Information items are independent from each other

Bayesian Merging

- **Idea** : there is a unique probability distribution capturing the behaviour of sources.
- **Data**:
 - x_i : observation of the value of quantity x by source i .
 - $P(x_1 \text{ and } x_2 | x)$ information about source behaviour
 - $P(x)$ prior information about the value of x

- $$P(x | x_1 \text{ and } x_2) = \frac{P(x_1 \text{ and } x_2 | x) \cdot P(x)}{\sum_{x'} P(x_1 \text{ and } x_2 | x') \cdot P(x')}$$

- (requires a lot of data)

« Idiot Bayes »

- Usual assumption: precise observations x_1 and x_2 are conditionally independent with respect to x .

$$P(x_1 | x) \cdot P(x_2 | x) \cdot P(x)$$

- $P(x | x_1 \text{ and } x_2) = \frac{P(x_1 | x) \cdot P(x_2 | x) \cdot P(x)}{\sum_{x'} P(x_1 | x') \cdot P(x_2 | x') \cdot P(x')}$
 - Independence assumption often unrealistic
 - Conjunctive product-based combination rule similar to possibilistic merging, if we let $P(x_i | x) = \pi_i(x)$
- *A likelihood function is an example of a possibility distribution*

- *What if no prior information?* Bayesians use Laplace principle: A uniform prior

$$P(x_1 | x) \cdot P(x_2 | x)$$

- $P(x | x_1 \text{ and } x_2) = \frac{P(x_1 | x) \cdot P(x_2 | x)}{\sum_{x'} P(x_1 | x') \cdot P(x_2 | x')}$

- **Too strong** : merging likelihood functions should yield a likelihood function.

$$P(x_1 | x) \cdot P(x_2 | x)$$

- $\pi(x) = \frac{P(x_1 | x) \cdot P(x_2 | x)}{\sup_{x'} P(x_1 | x') \cdot P(x_2 | x')}$ possibilistic merging

Possibilistic merging with prior information

- *Bayes theorem* :

$$\pi(u_1, u_2 | u) * \pi_x(u) = \pi_x(u | u_1, u_2) * \pi(u_1, u_2).$$

- $\pi_x(u)$ a priori information about x (uniform = ignorance)
- $\pi(u_1, u_2 | u)$: results from a merging operation F
- $\pi(u_1, u_2) = \sup_{u \in U} \pi(u_1, u_2 | u) \cdot \pi_x(u)$.

- If operation F is product:

$$\pi(u) = \frac{\pi(u_1 | u) \cdot \pi(u_2 | u) \cdot \pi_x(u)}{\sup_{u'} \pi(u_1 | u') \cdot \pi(x_2 | u') \cdot \pi_x(u)}$$

- Similar to probabilistic Bayes **but** more degrees of freedom

MERGING BELIEF FUNCTIONS

- Problem :
- source $i \rightarrow (\mathcal{F}^i, m_i)$ with $\sum_{A \in \mathcal{F}^i} m_i(A) = 1$
- Dempster rule of combination : *an associative scheme* generalising Dempster conditioning
 - Step 1 : $m_{\cap}(C) = \sum_{A \cap B = C} m_1(A) \cdot m_2(B)$
Independent random set intersection
 - Step 2 : $m^*(C) = m_{\cap}(C) / (1 - m_{\cap}(\emptyset))$
renormalisation $m_{\cap}(\emptyset)$ evaluates conflict ; it is eliminated.

Example : $S = \{a, b, c, d\}$

m_2	$\{c\}$	$\{b, c, d\}$	S
m_1	0.2	0.7	0.1
$\{b\}$	\emptyset	$\{b\}$	$\{b\}$
0.3	0.06	0.21	0.03
$\{a, b, c\}$	$\{c\}$	$\{b, c\}$	$\{a, b, c\}$
0.5	0.1	0.35	0.05
S	$\{c\}$	$\{b, c, d\}$	S
0.2	0.04	0.14	0.02

$$m_{\cap}(\{b\}) = 0.21 + 0.03 = 0.24 ; m_{\cap}(\{c\}) = 0.1 + 0.04 = 0.15$$

$$m_{\cap}(S) = 0.02 ; m_{\cap}(\emptyset) = 0.06$$

Disjunctive merging of belief functions

$$m_{\cup}(C) = \sum_{C : A \cup B = C} m_1(A).m_2(B)$$

- *Union of independent random sets.*
- More imprecise than conjunctive merging, even normalised.
- Moreover $\text{Bel}_{\cup}(A) = \text{Bel}_1(A).\text{Bel}_2(A)$
 - Disjunctively combining two probability distributions yields a random set.
 - *Belief functions are closed via product and convex sum.*
 - If conflict is too strong, normalized conjunctive merging provides arbitrary results and should be avoided: use another scheme like disjunctive merging.

Conjunctive merging with disjunctive conflict management

1. Conflict is ignorance

$$\begin{aligned} - \quad m_{\cap\delta}(C) &= \sum_{A \cap B = C} m_1(A).m_2(B) \text{ if } C \neq \emptyset, S \\ - \quad m_{\cap\delta}(S) &= \sum_{A \cap B = \emptyset} m_1(A).m_2(B) + m_1(S).m_2(S) \\ &= m_{\cap}(\emptyset) + m_{\cap}(S) \end{aligned}$$

2. Adaptive rule: for $C \neq \emptyset$

$$m_{\cap\delta}(C) = \sum_{A \cap B = C} m_1(A).m_2(B) + \sum_{\substack{A \cup B = C \\ A \cap B = \emptyset}} m_1(A).m_2(B)$$

These rules are not associative.

Compromise merging

- **Convex combination:** generalisation of the probabilistic merging rule
- $m_\alpha(A) = \alpha \cdot m_1(A) + (1 - \alpha) m_2(A)$
 - α = relative reliability of source 1 versus source 2
- **Example :** discounting an unreliable belief function with reliability α close to 1: combine m_1 with the void belief function $m_2(S) = 1$: then
 - $m_\alpha(A) = \alpha \cdot m_1(A)$ if $A \neq S$
 - $m_\alpha(A) = \alpha \cdot m_1(S) + (1 - \alpha)$

CONCLUSION: Belief construction for an agent

1. **Perception** : collecting evidence tainted with uncertainty
 2. **Merging** : Combining new evidence with current one so as to lay bare an (incomplete) description of the current situation considered as true.
 3. **Plausible inference** :Forming beliefs by applying background knowledge to current evidence
- This scheme can be applied in various settings encompassed by imprecise probability, but
 - classical logic is too poor : need conditional events. Non-monotonic reasoning a la Lehmann (or qualitative possibilistic logic) is minimal requirement for step 3.
 - Bayesian probability is too rich : ever complete and consistent.
 - Shafer-Smets or possibility theory is useful for merging uncertain evidence(step 2)