

Introduction to Machine Learning

Week 1

Lecture 2: Random Variables

Prof. Subrahmanya Swamy Peruru
Department of Electrical Engineering
IIT Kanpur

Recap

Probability ($\Omega, \mathcal{F}, P$)

Sample Space $\Omega = \{H, T\}$

Collection of Events $\mathcal{F} = \{\emptyset, \Omega, \{H\}, \{T\}\}$

Probability Measure

$P : \mathcal{F} \rightarrow [0, 1]$

- $P(\emptyset) = 0$
- $P(\Omega) = 1$
- $P(\{H\}) = 0.5$
- $P(\{T\}) = 0.5$

Axioms

- Normalization: $P(\Omega) = 1$
- Countable Additivity

Properties

Complement: $P(A^c) = 1 - P(A)$

Inclusion-Exclusion: $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

Union Bound: $P(A \cup B) \leq P(A) + P(B)$

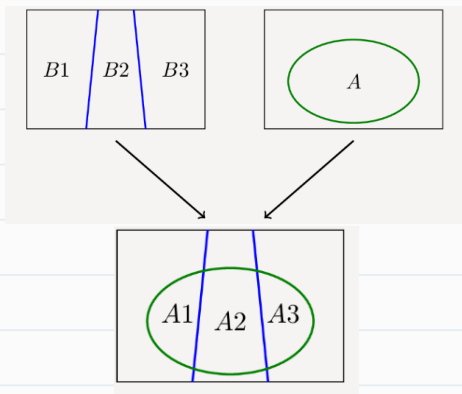
Recap

Conditional Probability

A given B: $P(A | B) = \frac{P(A \cap B)}{P(B)}$

Independence: $P(A \cap B) = P(A) P(B)$
(or)
 $P(A | B) = P(A)$

LoTP: $P(A) = P(A|B_1)P(B_1) + P(A|B_2)P(B_2) + P(A|B_3)P(B_3)$



Recap

Bayes Rule

- Consider our two-coin experiment discussed before

Question: Given the outcome is a head, what is the probability that coin 1 was chosen?

$$\begin{aligned} P(\text{Coin 1} \mid \text{Head}) &= \frac{P(\text{Coin 1} \cap \text{Head})}{P(\text{Head})} \\ &= \frac{P(\text{Head} \mid \text{Coin 1}) P(\text{Coin 1})}{P(\text{Head} \mid \text{Coin 1}) P(\text{Coin 1}) + P(\text{Head} \mid \text{Coin 2}) P(\text{Coin 2})} \quad (\text{LoTP}) \end{aligned}$$

Prior: $P(\text{Coin 1})$

Likelihood: $P(\text{Head} \mid \text{Coin 1})$

Posterior: $P(\text{Coin 1} \mid \text{Head})$

Recap

Conditional Independence



- When you toss a coin twice
- Outcome of **first toss is independent of the second** toss
 - H_1 : Head in *first* toss
 - H_2 : Head in *second* toss
 - $P(H_1 \cap H_2) = P(H_1)P(H_2)$

Definition

Let $P(C) > 0$.

Then A, B are **conditionally independent given C** if

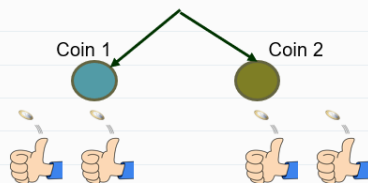
$$P(A \cap B | C) = P(A | C) P(B | C)$$

Another Experiment:

Suppose there are two coins:

- **Coin 1** lands heads with probability **0.9**
- **Coin 2** lands heads with probability **0.1**

Select one of the two coins uniformly at random



Toss the selected coin twice

Exercise:

Are H_1 and H_2 independent in this new experiment?

Are H_1 and H_2 conditionally independent given that Coin 1 was chosen?

Lecture Content

1

Random Variables

- Discrete RV
- Continuous RV
- CDF, PMF, PDF

2

Expectation

- Mean
- Variance
- Linearity of Expectation

3

Multiple Random Variables

- Joint, Marginal, Conditional Dist.
- Independence
- Correlation

Random variable

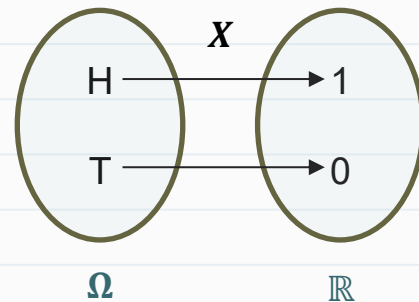
- A random variable **maps** the **outcomes** of random experiment **to real numbers**

$$X: \Omega \rightarrow \mathbb{R}$$

- We often write $X(\omega)$ as the value of the RV when outcome ω occurs.

$$X(H) = 1$$

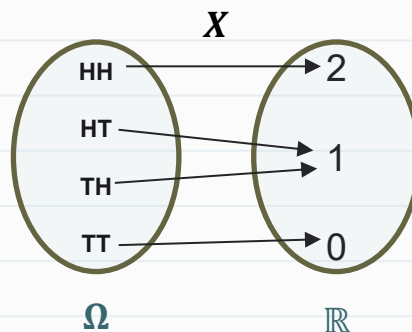
$$X(T) = 0$$



Example:

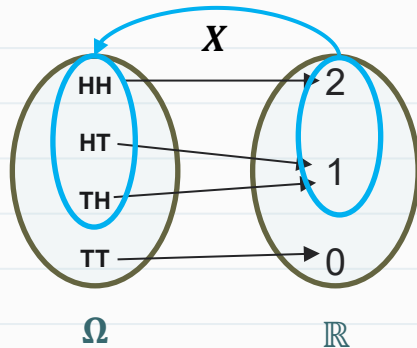
Experiment: Toss a coin twice
Random Variable: Number of Heads

$$\Omega = \{HH, HT, TH, TT\}$$



How to compute probabilities of RV ?

$X = \# \text{ Heads when a fair coin tossed twice}$



$$\mathbb{P}(X \in \{1, 2\}) = \mathbb{P}(\{HH, HT, TH\}) = \frac{3}{4}$$

In General:

$$\mathbb{P}(X \in A) = \mathbb{P}(\omega \mid X(\omega) \in A)$$

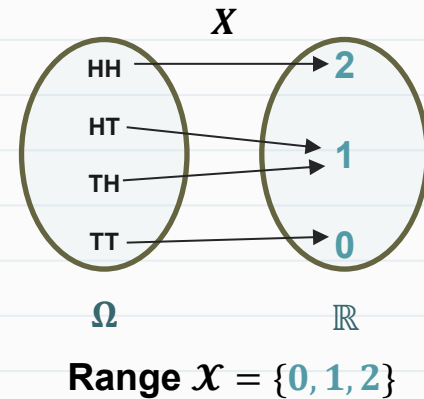
Types of RV:

- Discrete
- Continuous

Discrete RV

The **range** (support) of X is the set of values it can take:

$$\mathcal{X} = \{X(\omega): \omega \in \Omega\}$$



Discrete RV: range is countable (finite or countably infinite).

PMF: probability mass function

If X is **discrete**, define the **PMF**:

$$p_X(x) = \mathbb{P}(X = x), \quad x \in \mathcal{X}.$$

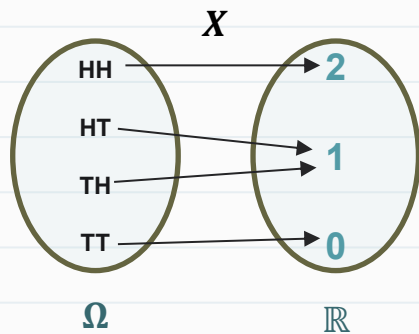
Properties:

$$p_X(x) \geq 0, \quad \sum_{x \in \mathcal{X}} p_X(x) = 1.$$

For any event $\{X \in A\}$ with $A \subseteq \mathcal{X}$:

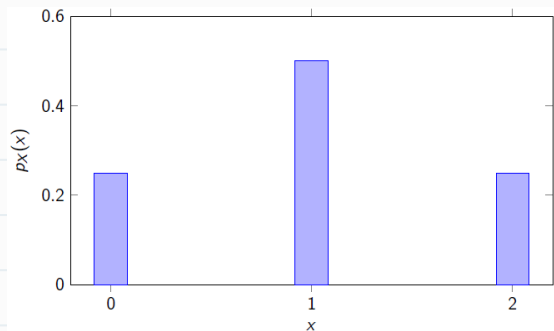
$$\mathbb{P}(X \in A) = \sum_{x \in A} p_X(x).$$

Example: PMF



Range $\mathcal{X} = \{0, 1, 2\}$

$$p_X(0) = \mathbb{P}(TT) = \frac{1}{4}, \quad p_X(1) = \mathbb{P}(HT \text{ or } TH) = \frac{1}{2}, \quad p_X(2) = \mathbb{P}(HH) = \frac{1}{4}$$



Verify: $\sum_{x \in \mathcal{X}} p_X(x) = 1$

$$\frac{1}{4} + \frac{1}{2} + \frac{1}{4} = 1$$

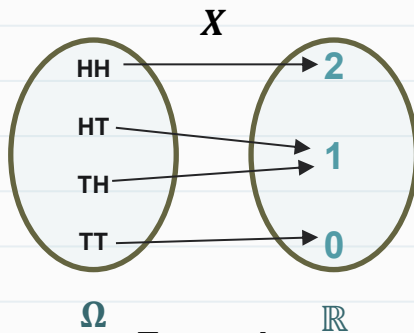
PMF is a function on discrete points; total mass sums to 1

CDF: Cumulative Distribution Function

PMF: Only for Discrete RV,

CDF: For any RV

$$F_X(\mathbf{x}) = \mathbb{P}(\mathbf{X} \leq \mathbf{x})$$



Example

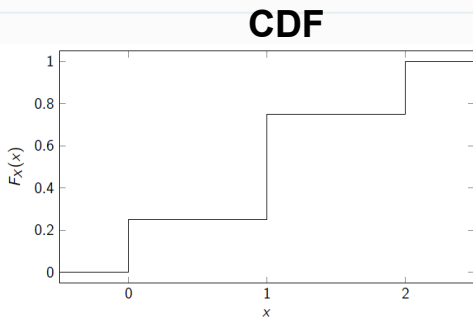
- $F_X(-1) = \mathbb{P}(X \leq -1) = 0$

- $F_X(0) = \mathbb{P}(X \leq 0) = \frac{1}{4}$

- $F_X(1) = \mathbb{P}(X \leq 1) = \frac{3}{4}$

- $F_X(2) = \mathbb{P}(X \leq 2) = 1$

- $F_X(3) = \mathbb{P}(X \leq 3) = 1$

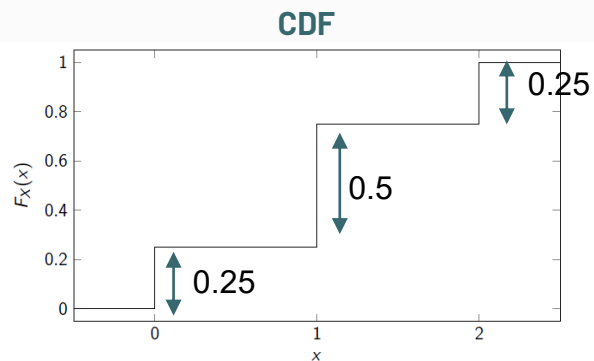
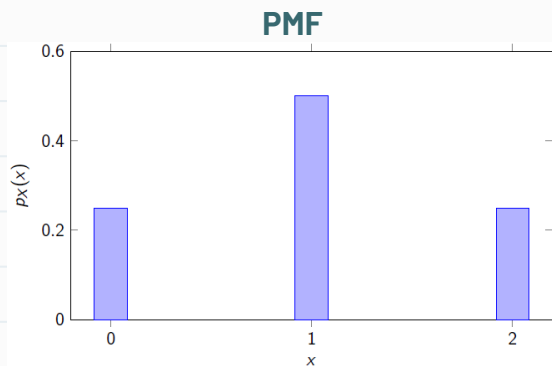


- $F_X(x)$ is **non-decreasing** in x .

- $\lim_{x \rightarrow -\infty} F_X(x) = 0$, $\lim_{x \rightarrow +\infty} F_X(x) = 1$.

- For any $a < b$: $\mathbb{P}(a < \mathbf{X} \leq \mathbf{b}) = F_X(\mathbf{b}) - F_X(\mathbf{a})$.

Discrete RVs: PMF and CDF



Discrete CDF is a **step function**

Jumps equal the PMF masses

$$\mathbb{P}(X \leq x) = F_X(x) = \sum_{y=-\infty}^x p_X(y)$$

Moving to Continuous RVs

Suppose X takes one of the three values in $\{2, 2.25, 2.5\}$ uniformly at random

Discrete

$$\mathbb{P}(X = 2) = \frac{1}{3}$$

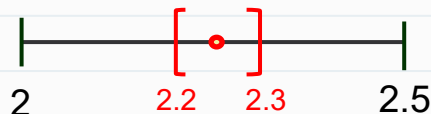
$$\mathbb{P}(X = 2.25) = \frac{1}{3}$$

$$\mathbb{P}(X = 2.5) = \frac{1}{3}$$

PMF

Suppose X takes a real number in the interval $[2, 2.5]$ uniformly at random

What is $\mathbb{P}(X = 2.25) = 0$



What is $\mathbb{P}(X \in [2.2, 2.3])$ → Can be answered using PDF

Continuous RVs: PDF

PDF is the derivative of CDF:

The probability density function (PDF) is given by

$$f_X(x) = \frac{dF_X(x)}{dx}, \quad (\text{wherever } F_X(x) \text{ is differentiable})$$

Properties:

$$f_X(x) \geq 0$$

$$\mathbb{P}(a \leq X \leq b) = \int_a^b f_X(t) dt$$

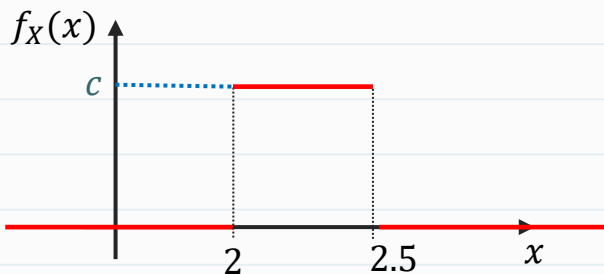
$$\int_{-\infty}^{\infty} f_X(t) dt = 1$$

CDF is an integral of PDF:

$$\mathbb{P}(X \leq x) = F_X(x) = \int_{-\infty}^x f_X(t) dt$$

Example: Continuous RV

Example: $X \sim \text{Unif}[2, 2.5]$



$$\int_{-\infty}^{+\infty} f_X(x) dx = 1 \Rightarrow \int_2^{2.5} c dx = 1$$

$$\Rightarrow c = 2$$

Important Observation:

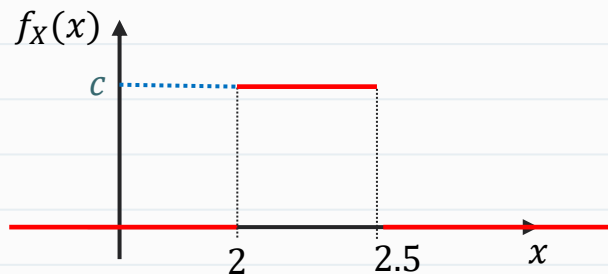
- The value of a PDF can be greater than 1.
- This is because the PDF itself is **not a probability**.
- **Probabilities are obtained by integrating the PDF.**

Compute $\mathbb{P}(2.25 \leq X \leq 2.5)$

$$\int_{2.25}^{2.5} f_X(x) dx = \int_{2.25}^{2.5} 2 dx = 2 \times (2.5 - 2.25) = 0.5$$

Example: Continuous RV

Exercise: Compute **CDF** of $X \sim \text{Unif}[2, 2.5]$



Integrate the PDF

$$F_X(x) = \mathbb{P}(X \leq x) = \int_{-\infty}^x f_X(t) dt$$

Expectation of a random variable

For discrete X :

$$\mu(X) = \mathbb{E}[X] = \sum_x x p_X(x)$$

Discrete $\rightarrow$ Continuous:

PMF $\rightarrow$ PDF

Sum $\rightarrow$ Integral

For continuous X :

$$\mu(X) = \mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

Example (Bernoulli RV): Find $E[X]$

when the PMF is given by

$$p_X(1) = p$$

$$p_X(0) = 1 - p$$

$$E[X] = 1 \times p + 0 \times (1 - p) = p$$

Function of a RV: $Y = g(X)$ (For example, $Y = X^2$)

Exercise: For the above random variable X , **Find $E[X^3 + 2X + 3]$**

How to find $\mathbb{E}[Y]$?

Approach 1: Compute the PMF $p_Y(y)$ and then compute $\mathbb{E}[Y] = \sum_y y p_Y(y)$

Approach 2 (Easier): Directly compute $\mathbb{E}[Y] = \mathbb{E}[g(X)] = \sum_x g(x) p_X(x)$

Note: Expectation exists only when the absolute-value integrals/sums are finite

Expectation is Linear

For any $a, b \in \mathbb{R}$:

$$\mathbb{E}[aX + b] = a \mathbb{E}[X] + b$$

For any set of RVs $X_1, X_2, \dots, X_n$:

$$\mathbb{E}[X_1 + X_2 + \dots + X_n] = \mathbb{E}[X_1] + \mathbb{E}[X_2] + \dots + \mathbb{E}[X_n]$$

Example: $X \sim \text{Binomial}(n, p)$. Find $\mathbb{E}[X]$.

- Consider a coin whose **probability of heads is p**
- Let X be the **number of heads** observed in **n independent tosses of the coin**.

Using Standard Approach:

- Start with the Binomial PMF $p_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$
- Compute $\mathbb{E}[X] = \sum_x x p_X(x)$

Using Linearity of Expectation:

- Let $X_1, X_2, \dots, X_n$ be *Bernoulli*(p) RVs
- taking a value 1/0 depending on whether the corresponding toss is a head/tail
- $X = X_1 + X_2 + \dots + X_n$
- $\mathbb{E}[X] = \mathbb{E}[X_1] + \mathbb{E}[X_2] + \dots + \mathbb{E}[X_n] = p + p + \dots + p = np$

Variance and standard deviation

Variance: $\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$

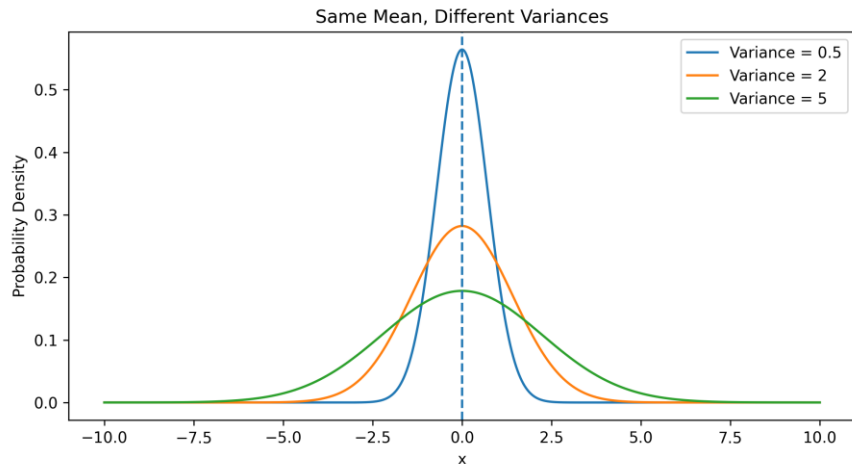
Variance measures how spread out the values are around the mean

Standard deviation: $\sigma_X = \sqrt{\text{Var}(X)}$

Example: Gaussian Distribution

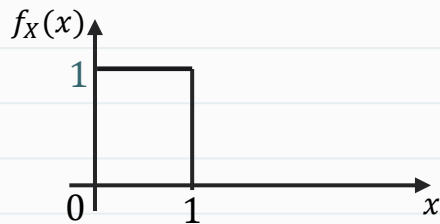
$X \sim \mathcal{N}(\mu, \sigma^2)$:

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$



Examples : $\mathbb{E}[X]$ and $\text{Var}(X)$

Uniform RV $X \sim \text{Unif}(0,1)$



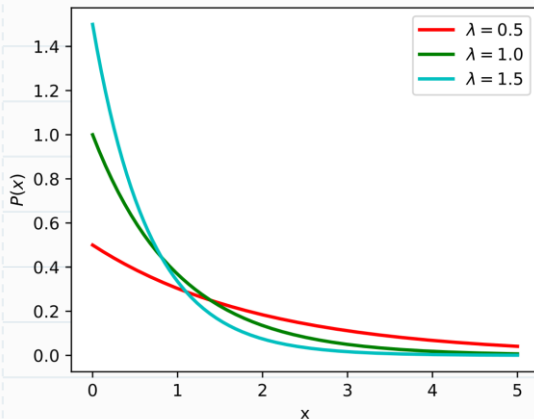
$$\circ \mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x dx = \frac{1}{2}$$

$$\circ \mathbb{E}[X^2] = \int_0^1 x^2 dx = \frac{1}{3}$$

$$\circ \text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}$$

Exercise: Solve this problem for $X \sim \text{Unif}(a,b)$

Exponential RV $X \sim \text{Exp}(\lambda)$



$$\circ f_X(x) = \lambda e^{-\lambda x}, \quad x \geq 0$$

$$\circ \mathbb{E}[X] = \int_0^{\infty} x \lambda e^{-\lambda x} dx = \frac{1}{\lambda}$$

$$\circ \text{Var}(X) = \frac{1}{\lambda^2}$$

Multiple RVs: Joint PMF (discrete)

Suppose I want to characterize two random variables (X, Y)

Is it sufficient to know their PMFS: $p_X(x), p_Y(y)$?

Example: Predict rain or not based on

$X = \text{Temperature}$

$Y = \text{Humidity}$

NO

For discrete X, Y , **joint PMF**:

$$p_{X,Y}(x, y) = \mathbb{P}(X = x, Y = y).$$

$$\sum_{x,y} p_{XY}(x, y) = 1$$

Example:

$X = \text{Cointoss } \{0, 1\},$

$Y = 3 - \text{Faced die } \{1, 2, 3\}$

	$y = 0$	$y = 1$	$y = 2$
$x = 0$	$p_{XY}(0, 0)$	$p_{XY}(0, 1)$	$p_{XY}(0, 2)$
$x = 1$	$p_{XY}(1, 0)$	$p_{XY}(1, 1)$	$p_{XY}(1, 2)$

Marginal PMF

Marginals: $p_X(x)$, $p_Y(y)$

Relation to Joint PMF:

$$p_X(x) = \sum_y p_{X,Y}(x, y), \quad p_Y(y) = \sum_x p_{X,Y}(x, y).$$

Marginal Probability Example:

Suppose (X, Y) has joint PMF:

	$y = 0$	$y = 1$	$y = 2$
$x = 0$	0.10	0.10	0.05
$x = 1$	0.20	0.25	0.10
$x = 2$	0.05	0.10	0.05

Check total probability: sum of all entries = 1.

Compute marginals:

$$p_X(1) = 0.20 + 0.25 + 0.10 = 0.55,$$

$$p_Y(1) = 0.10 + 0.25 + 0.10 = 0.45.$$

Can we find the joint PMF from the marginals PMFs ?

No (unless they are independent)

Independence of RVs

Discrete case: X and Y are independent if

$$p_{X,Y}(x,y) = p_X(x)p_Y(y) \quad \forall \mathbf{x}, \mathbf{y}$$

Example

	$y = 0$	$y = 1$	$y = 2$
$x = 0$	0.10	0.10	0.05
$x = 1$	0.20	0.25	0.10
$x = 2$	0.05	0.10	0.05

$$p_{X,Y}(1,1) = 0.25.$$

$$p_X(1) = 0.20 + 0.25 + 0.10 = 0.55,$$

$$p_Y(1) = 0.10 + 0.25 + 0.10 = 0.45.$$

But

$$p_X(1)p_Y(1) = 0.55 \times 0.45 = 0.2475 \neq 0.25.$$

So X and Y are **not independent**.

Conditional PMF (discrete)

If $p_Y(y) > 0$, define:

$$p_{X|Y}(x|y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}.$$

Using the table:

$$p_{X|Y}(1|1) = \frac{p_{X,Y}(1,1)}{p_Y(1)} = \frac{0.25}{0.45} = \frac{5}{9} \approx 0.556.$$

Joint PMF

	$y = 0$	$y = 1$	$y = 2$
$x = 0$	0.10	0.10	0.05
$x = 1$	0.20	0.25	0.10
$x = 2$	0.05	0.10	0.05

Expectation with two RVs

$$\mathbb{E}[X] = \sum_x x p_X(x), \quad \mathbb{E}[Y] = \sum_y y p_Y(y).$$

A function of two RVs $g(\mathbf{X}, \mathbf{Y})$ (Example: $g(X, Y) = XY$)

$$\mathbb{E}[g(\mathbf{X}, \mathbf{Y})] = \sum_x \sum_y g(x, y) p_{\mathbf{X}, \mathbf{Y}}(\mathbf{x}, \mathbf{y}).$$

Useful identity (if independent):

$$X \perp Y \quad \Rightarrow \quad \mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y].$$

Recall: **(Linearity of Expectation even if X, Y are not independent)**

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y].$$

Conditional Expectation

$$\mathbb{E}[Y|X = x] = \sum_y y p_{Y|X}(y|x)$$

Example: Compute $\mathbb{E}[Y | X = 1]$

$$p_{Y|X}(y|1) = \frac{p_{X,Y}(1,y)}{p_X(1)} = \frac{p_{X,Y}(1,y)}{0.55}.$$

So:

Joint PMF

	$y = 0$	$y = 1$	$y = 2$
$x = 0$	0.10	0.10	0.05
$x = 1$	0.20	0.25	0.10
$x = 2$	0.05	0.10	0.05

$$\begin{aligned}\mathbb{E}[Y | X = 1] &= \sum_y y p_{Y|X}(y|1) \\ &= 0 \times p_{Y|X}(0 | 1) + 1 \times p_{Y|X}(1 | 1) + 2 \times p_{Y|X}(2 | 1) \\ &= 0 \times \frac{0.20}{0.55} + 1 \times \frac{0.25}{0.55} + 2 \times \frac{0.10}{0.55} \\ &= 0.82\end{aligned}$$

Covariance and correlation

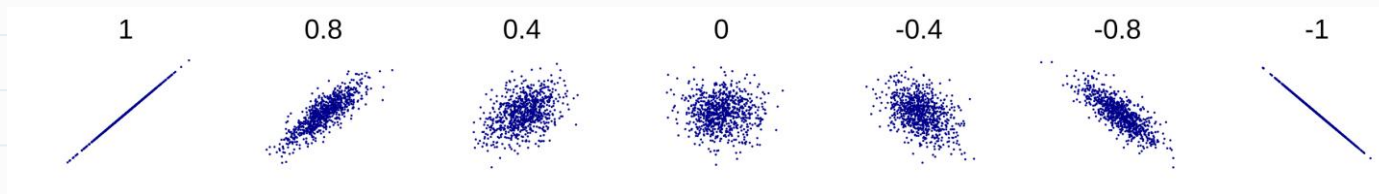
Covariance:

$$\text{Cov}(X, Y) = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y].$$

Correlation Coefficient:

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}} \in [-1, 1].$$

Several sets of (x, y) points, with the correlation coefficient of x and y for each set.



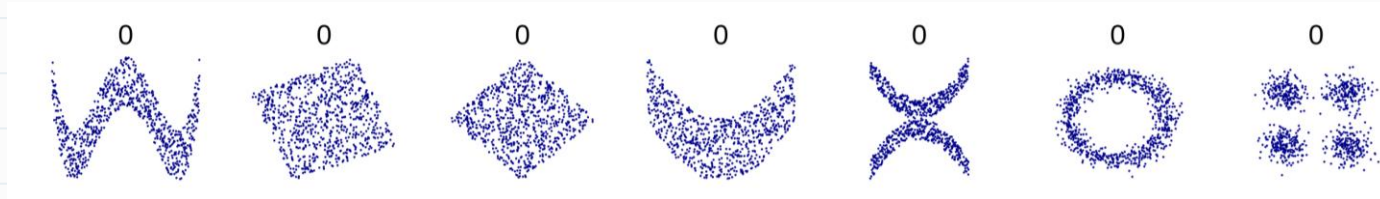
The correlation captures

- the presence of a **linear relationship**
- and how **noisy** that relationship is

Uncorrelated RVs: X and Y are uncorrelated if $\text{Cov}(X, Y) = 0$ (or equivalently) $\rho_{X,Y} = 0$

Zero Correlation?

Correlation does not capture non-linear relationships



Several sets of (x, y) points, with the correlation coefficient of x and y for each set.

Independence $\Rightarrow$ zero covariance, but **not conversely**

Independent $\Rightarrow$ Uncorrelated

Uncorrelated $\nRightarrow$ Independent

● Example: uncorrelated but dependent

● Let $\mathbf{X} \sim \mathbf{Unif}(-1, 1)$ and define $\mathbf{Y} = \mathbf{X}^2$.

● Then $\mathbb{E}[\mathbf{X}] = 0$ by symmetry.

● Compute:

$$\text{Cov}(\mathbf{X}, \mathbf{Y}) = \mathbb{E}[\mathbf{XY}] - \mathbb{E}[\mathbf{X}]\mathbb{E}[\mathbf{Y}] = \mathbb{E}[\mathbf{X}^3] - 0.$$

● But $\mathbf{X}^3$ is an odd function and $\mathbf{X}$ is symmetric on $[-1, 1]$, so $\mathbb{E}[\mathbf{X}^3] = 0$.

● Thus **Cov(X, Y) = 0** (uncorrelated), **yet Y is a deterministic function of X** $\Rightarrow$ dependent.

References

- Chapters 3-5 of the book **H. Pishro-Nik, "Introduction to probability, statistics, and random processes"**, available at <https://www.probabilitycourse.com>, Kappa Research LLC, 2014.

Thank You